Fiscal foresight and the effects of government spending: It's all in the monetary-fiscal mix¹

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^{1*} Views expressed are those of the authors and do not reflect official positions of respective institutions of affiliation.

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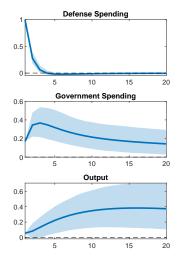
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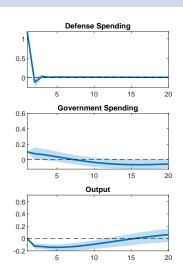
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- This evidence favours a neoclassical view over a Keynesian one

Empirical evidence by subsamples

Ramey (2011) defence spending shocks



Great Inflation (1960q1-1979q2)



Great Moderation (1984q1-2007q2)

Why?

We extend the analysis taking monetary-fiscal interactions into account

Great Moderation ⇒ Monetary regime (M)

central bank sets interest rates government adjusts deficits to stabilize real debt ⇒ Ricardian equivalence holds, no wealth effects on debt

Great Inflation \Rightarrow Fiscal regime (F)

government sets deficits, central bank accommodates rates to let inflation stabilize real debt

 \Rightarrow fiscal theory at work, no Ricardian equivalence, wealth effects on debt

(e.g., Bianchi and Melosi, 2017, 2014; Bianchi, 2012; Chung et al., 2007; Sims, 2011; Davig and Leeper, 2007, 2011)

Does fiscal foresight lead fiscal shocks to have different effects under regime M or F?

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- Explore the issue of non-fundamentalness

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- Conditional on the regime, shocks are fundamental: Granger causality tests ⇒ no evidence that shocks could have been forecasted

A simple New Keynesian model with fiscal policy block

Beck-Friis and Willems (2017)

$$\begin{split} \hat{y}_t - \alpha_1 \tilde{g}_t &= \mathbb{E}_t \hat{y}_{t+1} - \alpha_1 \mathbb{E}_t \tilde{g}_{t+1} - \alpha_2 [\hat{\imath}_t - \mathbb{E}_t \hat{\pi}_{t+1}] \qquad \text{(Euler)} \\ \hat{\pi}_t &= \beta \mathbb{E}_t [\hat{\pi}_{t+1}] + \kappa \alpha_3 \hat{y}_t - \kappa \alpha_4 \tilde{g}_t \qquad \text{(Phillips curve)} \\ \hat{\imath}_t &= \phi \hat{\pi}_t \qquad \qquad \text{(Taylor rule)} \\ \tilde{b}_t &= \frac{1}{\beta} \tilde{b}_{t-1} - \frac{1}{\beta} (\tilde{\tau}_t - \tilde{g}_t) - \frac{1}{\beta} \frac{b}{y} \hat{\pi}_t + \frac{b}{y} \hat{\imath}_t \qquad \text{(Government b.c.)} \\ \tilde{\tau}_t &= \psi \tilde{b}_{t-1} + \varepsilon_t^{\tau} \qquad \qquad \text{(Tax rule)} \\ \tilde{g}_t &= \rho \tilde{g}_{t-1} + \varepsilon_t^{g} \qquad \text{(Government spending rule)} \end{split}$$

Monetary and fiscal parametrization

Monetary regime

- active monetary policy: $\phi > 1$
- passive fiscal policy: $\psi > 1 \beta$

Fiscal regime

- ullet passive monetary policy: $\phi < 1$
- active fiscal policy: $\psi < 1 \beta$

We calculate analytically both anticipated and unanticipated GS multipliers on output for different degree of anticipation under the two regimes

Monetary regime

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 - shifts of labor supply (Neoclassical negative wealth effect)
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- Unanticipated multipliers: Positive
 - shifts of labor supply (Neoclassical negative wealth effect)
 - shifts of labor demand (sticky prices)
- Anticipated multipliers : Negative
 - Two competing effects
 - 1. $\pi^e \uparrow \Rightarrow r \downarrow \Rightarrow c \uparrow \text{ (real interest rate channel)}$
 - 2. $c^e \downarrow \Rightarrow c \downarrow \text{ (negative wealth effect + consumption smoothing)}$
 - Potentially ambiguous response of current demand
 - For a standard calibration, demand and output decreases in the anticipation period
 - anticipated G shock lowers consumption, as it is fiscally backed

Effects on inflation and real debt

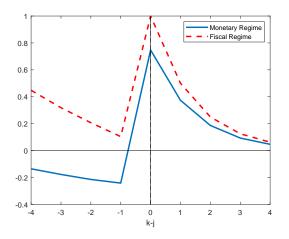
Fiscal regime

- Unanticipated multipliers: Positive
 - Nominal wealth effects (shut down under M)
 - $TM_F^y(0, k)$ newly-issued bonds increase net wealth for the households (equivalent to a debt-financed tax cut)
 - Keynesian and nominal wealth effects are both expansionary upon implementation (our calibration)
 - $GSM_F^y > GSM_M^y$

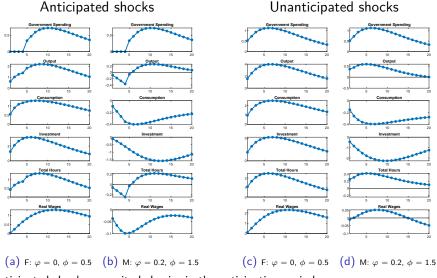
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 - $GSM_F^y > GSM_M^y$
- Anticipated multipliers: Positive
 - Same intuition: Keynesian effect + nominal wealth effect
 - Keynesian effect contractionary during anticipation
 - nominal wealth effect expansionary during anticipation
 - Nominal wealth effect dominates
 True houses of activity.
 - \Rightarrow two bursts of activity

Output response to an announced fiscal expansion

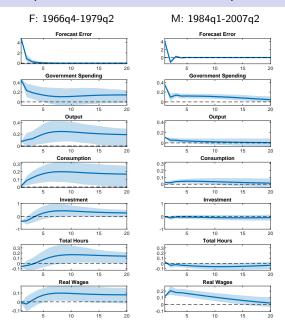


Smets and Wouters with 4-period anticipation



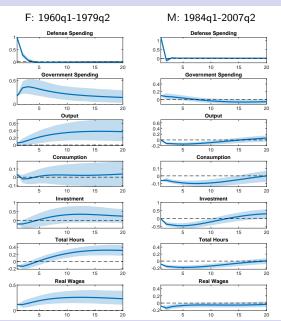
Anticipated shock: opposite behavior in the anticipation period; Unanticipated shock: same behavior (except for investment, consumption)

Empirical evidence: Unanticipated shocks



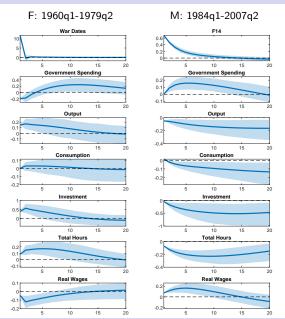
- Unanticipated shocks are identified as innovations to forecast errors using SPF data, Auerbach and Gorodnichenko (2012)
- no clear-cut differences between the two regimes

Additional empirical evidence: Anticipated shocks



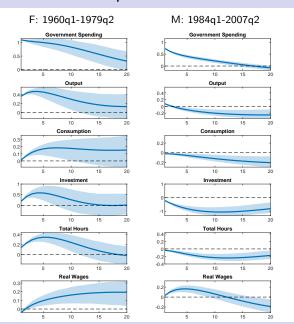
 Ramey (2011) defence spending shocks with larger VAR

Additional empirical evidence: Anticipated shocks



- F: Ramey and Shapiro (1998) war dates
- M: Forni and Gambetti (2016) ⇒ SPF forecast of future spending growth in the next four quarters F(1,4), ordered second in the VAR

Additional empirical evidence: Standard BP identification



- Blanchard and Perotti (2002) standard recursive identification with government spending ordered first (no distinction anticipated vs. unanticipated shocks)
- No difference with fiscal foresight VARs and theoretical results for anticipated shocks
- Under F: anticipated and unanticipated shocks are both expansionary, cannot tell apart
- Under M: fall in output thus supporting the idea that anticipation effects are the main transmission mechanism of fiscal shocks
- once controlling for the regime, shocks may become fundamental
- Tes

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- It could be (un)wise to anticipate future fiscal policies, depending on the regime in place.

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- Ambiguous effects on inflation
 - $Y \downarrow \Rightarrow \pi \downarrow$
 - $\pi^e \uparrow \Rightarrow \pi \uparrow$
- Ambiguous effects on real debt
 - depends on the real interest rate (cost of servicing the debt)
 - in turn, it depends on inflation and how hawkish the central bank is

Fundamentalness: Ramey's (2011) Granger causality test

Granger-causality tests between the residual from

the Blanchard and Perotti (2002) VAR and Ramey and Shapiro (1998) war dates

	Full sample 1947q1-2008q4	F 1960q1-1979q2	M 1984q1-2007q2
4 lags			
Do war dates Granger-cause VAR shocks?	Yes (0.0004)	No (0.5056)	No (0.5785)
Do VAR shocks Granger-cause war dates?	No (0.4938)	No (0.3803)	No (0.2415)
2 lags			
Do war dates Granger-cause VAR shocks?	Yes (0.0069)	No (0.2946)	No (0.4523)
Do VAR shocks Granger-cause war dates?	No (0.4776)	No (0.1997)	No (0.6601)

- VAR shocks never Granger-cause war dates
- War dates Granger cause VAR shocks only in the full sample

return Blanchard-Perotti VAR

Fundamentalness: Ramey's (2011) Granger causality test

Granger causality test using SPF forecasts of future spending growth for one and four quarters ahead					
	Full sample 1968q4-2008q4	F 1968q4-1979q2	M 1984q1-2007q2		
2 lags					
Do one-quarter ahead professional forecasts Granger-cause VAR shocks?	Yes (0.0667)	No (0.6320)	No (0.1711)		
Do VAR shocks Granger-cause one-quarter ahead professional forecasts?	No (0.3618)	No (0.6059)	No (0.2488)		
Do four-quarter ahead professional forecasts Granger-cause VAR shocks?			No (0.6577)		
Do VAR shocks Granger-cause four-quarter ahead professional forecasts?			No (0.1462)		

- Non-fundamentalness present in the full sample, which does not distinguish
 M and F
- When well-defined monetary and/or fiscal regimes are considered, shocks become fundamental

Testing for fundamentalness: Forni and Gambetti (2016)

Orthogonality test - Regime M					
	1 lag	2 lags	3 lags		
f(0)	0.85	0.78	0.81		
f(1)	0.17	0.25	0.35		
f(2)	0.75	0.14	0.23		
f(3)	0.99	0.93	0.04		
f(4)	0.87	0.51	0.50		
f(0) to $f(4)$	0.59	0.13	0.13		
F(1,4)	0.55	0.81	0.70		

- Except for one case with 3 lags, there is always evidence of fundamentalness
- And this is true even if one considers a smaller (4-variables) VAR