



Optimal reinsurance problem under fixed cost and exponential preferences

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Abstract: We investigate an optimal reinsurance problem for an insurance company taking into account subscription costs: that is, a constant fixed cost is paid when the reinsurance contract is signed. Differently from the classical reinsurance problem, where the insurer has to choose an optimal retention level according to some given criterion, in this paper the insurer needs to optimally choose both the starting time of the reinsurance contract and the retention level to apply. The criterion is the maximization of the insurer's expected utility of terminal wealth. This leads to a mixed optimal control/optimal stopping time problem, which is solved by a two-step procedure: first considering the pure-reinsurance stochastic control problem and next discussing a time-inhomogeneous optimal stopping problem with discontinuous reward. Using the classical Cramér-Lundberg approximation risk model, we prove that the optimal strategy is deterministic and depends on the model parameters. In particular, we show that there exists a maximum fixed cost that the insurer is willing to pay for the contract activation. Finally, we provide some economical interpretations and numerical simulations.

Keywords: Optimal Reinsurance; Mixed Control Problem; Optimal Stopping; Transaction Cost.

MSC: 93E20, 91B30, 60G40, 60J60.

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1. Introduction

Insurance business requires the transfer of risks from the policyholders to the insurer, who receives a risk premium as a reward. In some cases, it could be convenient to cede these risks to a third party, which is the reinsurance company. From the operational viewpoint, a risk-sharing agreement helps the insurer reducing unexpected losses, stabilizing operating results, increasing business capacity and so on. By means of a reinsurance treaty, the reinsurance company agrees to indemnify the primary insurer (cedent) against all or part of the losses which may occur under policies which the latter issued. The cedent will pay a reinsurance premium in exchange for this service. Roughly speaking, this is an insurance for insurers. When subscribing a reinsurance treaty, a natural question is to determine the (optimal) level of the retained losses. Optimal reinsurance problems have been intensively studied by many authors under different criteria, especially through expected utility maximization and ruin probability minimization, see for example [1], [2], [3], [4], [5] and references therein.

The main novelty of this article is that subscription costs are considered. Transaction costs represent the bureaucratic fixed costs necessary to run and manage an insurance company. The empirical impact of these costs on insurances choices has been highlighted in the actuarial literature, see for instance [6] and [7]. In practice, in the reinsurance

35 context, when the agreement is signed a fixed cost is usually paid in addition to the
36 reinsurance premium. This aspect has not been investigated by nearly all the studies,
37 except for [8] and [9]. In the former work the authors discussed the reinsurance prob-
38 lem subject to a fixed cost for buying reinsurance and a time delay in completing the
39 reinsurance transaction. They solved the problem considering a performance criterion
40 with linear current reward and showed that it is optimal to buy reinsurance when the
41 surplus lies in a bounded interval depending on the delay time. In the latter paper,
42 under the criterion of minimizing the ruin probability, the original problem is reduced to
43 a time-homogeneous optimal stopping problem. In particular, the authors show that the
44 fixed cost forces the insurer to postpone buying reinsurance until the surplus process
45 hits a certain level.

46 Hence the presence of a fixed cost is closely related to the possibility of postponing
47 the subscription of the reinsurance agreement. This, in turn, involves an optimal stopping
48 problem, which is attached to the optimal choice of the retention level, which is a well
49 known stochastic control problem. The novelty of our paper consists in considering this
50 mixed stochastic control problem under the criterion of maximizing the expected utility
51 of terminal wealth. The strategy of the insurance company consists of the retention level
52 of a proportional reinsurance and the subscription timing. When the contract is signed,
53 a given fixed cost is paid and the optimal retention level is applied. For the purpose
54 of mathematical tractability, we use a diffusion approximation to model the insurer's
55 surplus process (see [10]). The insurance company has exponential preferences and is
56 allowed to invest in a risk-less bond.

57 As already mentioned, this setup leads to a combined problem of optimal stopping
58 and stochastic control with finite horizon, which we will solve by a two-step procedure.
59 For theoretical studies on mixed control-stopping problems we refer to [11], [12] and
60 [13] among others. First, we provide the solution of the pure reinsurance problem (with
61 starting time equal to zero). Next, we discuss an optimal stopping time problem with
62 a suitable reward function depending on the value function of the pure reinsurance
63 problem. Differently to [8] and [9], the associated optimal stopping problem turns out to
64 be time-inhomogeneous and with discontinuous stopping reward with respect to the
65 time. We provide an explicit solution, also showing that the optimal stopping time is
66 deterministic. Moreover, we find that only two cases possible, depending on the model
67 parameters. When the fixed cost is greater than a suitable threshold (whose analytical
68 expression is available), the optimal choice is not to subscribe the reinsurance; otherwise,
69 the insurer immediately subscribes the contract.

70 A recent related research can be found in [14], where the problem of optimal
71 dividends and reinsurance is formulated as a mixed classical-impulse stochastic control
72 problem. The authors consider a fixed transaction cost when the dividends are paid out
73 and they solve the problem using the method of quasi-variational inequalities.

74 The paper is organized as follows. In Section 2, we describe the model and formulate
75 the problem as a mixed stochastic control problem, that is a problem which involves
76 both optimal control and stopping. In Section 3 we discuss the pure reinsurance problem
77 (without stopping) by solving the associated Halmilton-Jacobi-Bellman equation. Section
78 4 is devoted to the reduction of the original (mixed) problem to a suitable optimal
79 stopping problem, which is then investigated in Section 5. Here we provide a Verification
80 Theorem and we solve the associated variational inequality. In Section 6 we give the
81 explicit solution to the original problem and we discuss some economic implications of
82 our results. Finally, in Section 7 some numerical simulations are performed in order to
83 better understand the economic interpretation of our findings.

84 2. Problem formulation

85 2.1. Model formulation

86 Let $T > 0$ be a finite time horizon and assume that $(\Omega, \mathcal{F}, \mathbb{P}, \mathbb{F})$ is a complete proba-
87 bility space endowed with a filtration $\mathbb{F} \doteq \{\mathcal{F}_t\}_{t \in [0, T]}$ satisfying the usual conditions.

Let us denote by $R = \{R_t\}_{t \in [0, T]}$ the surplus process of an insurance company. There is a wide range of risk models in the actuarial literature, see for instance [15] and [10]. In the Cramér-Lundberg risk model the claims arrival times are described by the sequence of claims arrival times $\{T_n\}_{n \geq 1}$, with $T_n < T_{n+1}$ \mathbb{P} -almost everywhere $\forall n \geq 1$, while the corresponding claim sizes are given by $\{Z_n\}_{n \geq 1}$. In particular, the number of occurred claims up to time $t \geq 0$ is equal to

$$N_t = \sum_{n=1} \mathbb{1}_{\{T_n \leq t\}},$$

and it is assumed to be a Poisson process with constant intensity $\lambda > 0$, independent of the sequence $\{Z_n\}_{n \geq 1}$. Moreover, $\{Z_n\}_{n \geq 1}$ are independent and identically distributed random variables with common probability distribution function $F_Z(z)$, $z \in (0, +\infty)$, having finite first and second moments denoted by $\mu > 0$ and $\mu_2 > 0$, respectively. In this context the surplus process is given by

$$R_0 + ct - \sum_{n=1}^{N_t} Z_n, \quad R_0 > 0, \quad (2.1)$$

where R_0 is the initial capital and $c > 0$ denotes the gross risk premium rate. We can show that for any $t \geq 0$

$$\mathbb{E} \left[\sum_{n=1}^{N_t} Z_n \right] = \lambda \mu t \quad \text{and} \quad \text{var} \left[\sum_{n=1}^{N_t} Z_n \right] = \lambda \mu_2 t.$$

In this paper we use the diffusion approximation of the Cramér-Lundberg model (2.1), see for example [15]. Precisely, we assume that the surplus process follows this stochastic differential equation (SDE):

$$dR_t = p dt + \sigma_0 dW_t, \quad R_0 > 0,$$

88 where $W = \{W_t\}_{t \in [0, T]}$ is a standard Brownian motion, $\sigma_0 = \sqrt{\lambda \mu_2}$ and p denotes the
 89 insurer's net profit, that is $p = c - \mu \lambda$. In particular, under the expected value principle
 90 (see e.g. [10]) we have that $c = (1 + \theta_i) \mu \lambda$ and hence $p = \theta_i \mu \lambda$, with $\theta_i > 0$ representing
 91 the insurer's safety loading.

92

We allow the insurer to invest her surplus in a risk-free asset with constant interest rate $R > 0$:

$$dB_t = B_t R dt, \quad B_0 = 1,$$

hence the wealth process $X = \{X_t\}_{t \in [0, T]}$ evolves according to

$$dX_t = R X_t dt + p dt + \sigma_0 dW_t, \quad X_0 = R_0 > 0. \quad (2.2)$$

The explicit solution of the SDE (2.2) is given by the following equation:

$$X_t = R_0 e^{Rt} + \int_0^t e^{R(t-s)} p ds + \int_0^t e^{R(t-s)} \sigma_0 dW_s, \quad t \in [0, T]. \quad (2.3)$$

Now let τ denote an \mathbb{F} -stopping time. At time τ the insurer can subscribe a proportional reinsurance contract with retention level $u \in [0, 1]$, transferring part of her risks to the reinsurer. More precisely, u represents the percentage of retained losses, so that $u = 0$ means full reinsurance, while $u = 1$ is equivalent to no reinsurance. In order to buy a reinsurance agreement, the primary insurer pays a reinsurance premium $q(u) \geq 0$.

When the reinsurance contract is signed at time $t = 0$, the Cramér-Lundberg risk model (2.1) is replaced by the following equation:

$$R_0 + (c - q(u))t - \sum_{n=1}^{N_t} uZ_n, \quad R_0 > 0.$$

93 Under the expected value principle we have that $q(u) = (1 + \theta)(1 - u)\mu\lambda$, $u \in [0, 1]$,
 94 with the reinsurer's safety loading θ satisfying $\theta > \theta_i$ (preventing the insurer from
 95 gaining a risk-free profit).

Let us denote by $R^u = \{R_t^u\}_{t \in [0, T]}$ the reserve process in the Cramér-Lundberg approximation associated with a given reinsurance strategy $\{u_t\}_{t \in [0, T]}$ when the reinsurance contract is signed at time $t = 0$. Following [16], under the expected value principle, R^u follows

$$dR_t^u = (p - q + qu_t) dt + \sigma_0 u_t dW_t, \quad R_0^u = R_0, \quad (2.4)$$

where $q = \theta\lambda\mu$ denotes the reinsurer's net profit. We set $q > p$ (non-cheap reinsurance). The wealth process under the strategy $\{u_t\}_{t \in [0, T]}$ evolves according to this SDE:

$$dX_t^u = RX_t^u dt + dR_t^u, \quad X_0^u = R_0, \quad (2.5)$$

which admits this explicit representation:

$$X_t^u = R_0 e^{Rt} + \int_0^t e^{R(t-s)} (p - q + qu_s) ds + \int_0^t e^{R(t-s)} \sigma_0 u_s dW_s. \quad (2.6)$$

96 We assume that a constant fixed cost $K > 0$ is paid when the reinsurance contract is
 97 subscribed. The insurer decides when the reinsurance contract starts and which retention
 98 level is applied. Hence the insurer's strategy is a couple $\alpha = (\tau, \{u_t\}_{t \in [\tau, T]})$, with $\tau \leq T$.
 99 Let $H_t = I_{\{\tau \leq t\}}$ be the indicator process of the contract starting time. For $\tau < T$ \mathbb{P} -a.s.,
 100 the total wealth $X^\alpha = \{X_t^\alpha\}_{t \in [0, T]}$ associated with a given strategy α is given by

$$dX_t^\alpha = (1 - H_t)dX_t + H_t dX_t^u - K dH_t, \quad X_0^\alpha = R_0 > 0, \quad (2.7)$$

while on the event $\{\tau = T\}$ we have that

$$dX_t^\alpha = dX_t, \quad X_0^\alpha = R_0 > 0, \quad (2.8)$$

101 where X satisfies equation (2.2).

Equation (2.7) can be written more explicitly as

$$dX_t^\alpha = \begin{cases} dX_t, & t < \tau, \quad X_0 = R_0, \\ dX_t^u, & \tau < t \leq T, \quad X_\tau^u = X_\tau - K, \end{cases} \quad (2.9)$$

102 where X and X^u satisfy equations (2.2) and (2.5), respectively.

In our setting the null reinsurance corresponds to the choice $\tau = T$, \mathbb{P} -a.s., to which we associate the strategy $\alpha_{\text{null}} = (T, 1)$ and

$$X_t^{\alpha_{\text{null}}} = X_t, \quad t \in [0, T].$$

103 2.2. The utility maximization problem

The insurers' objective is to maximize the expected utility of the terminal wealth:

$$\sup_{\alpha \in \mathcal{A}} \mathbb{E}[U(X_T^\alpha)], \quad (2.10)$$

104 where $U : \mathbb{R} \rightarrow [0, +\infty)$ is the utility function representing the insurer's preferences and
 105 \mathcal{A} the class of admissible strategies (see Definition 1 below).

We focus on CARA (*Constant Absolute Risk Aversion*) utility functions, whose general expression is given by

$$U(x) = 1 - e^{-\eta x}, \quad x \in \mathbb{R},$$

106 where $\eta > 0$ is the risk-aversion parameter. This utility function is highly relevant in
107 economic science and particularly in insurance theory. Indeed, it is commonly used for
108 reinsurance problems (e.g. see [5] and references therein).

109

110 The optimization problem is a mixed optimal control problem. That is, the insurer's
111 controls involve the timing of the reinsurance contract subscription and the retention
112 level to apply.

113 **Definition 1** (Admissible strategies). We denote by \mathcal{A} the set of admissible strategies $\alpha =$
114 $(\tau, \{u_t\}_{t \in [\tau, T]})$, where τ is an \mathbb{F} -stopping time such that $\tau \leq T$ and $\{u_t\}_{t \in [\tau, T]}$ is an \mathbb{F} -
115 predictable process with values in $[0, 1]$. Let us observe that the null strategy $\alpha_{\text{null}} = (T, 1)$ is
116 included in \mathcal{A} . When we want to restrict the controls to the time interval $[t, T]$, we will use the
117 notation \mathcal{A}_t .

Proposition 1. Let $\alpha \in \mathcal{A}$, then

$$\mathbb{E}[e^{-\eta X_T^\alpha}] < +\infty.$$

Proof. Using equations (2.6) and (2.9), we have that

$$\begin{aligned} \mathbb{E}[e^{-\eta X_T^\alpha}] &= \mathbb{E}[e^{-\eta X_T} I_{\{\tau=T\}}] \\ &+ \mathbb{E}[e^{-\eta(X_\tau - K)e^{R(T-\tau)}} e^{-\eta \int_\tau^T e^{R(T-s)}(p-q+qu_s)ds} e^{-\eta \int_\tau^T e^{R(T-s)}\sigma_0 u_s dW_s} I_{\{\tau < T\}}]. \end{aligned}$$

Taking into account the expression (2.3) we get

$$\begin{aligned} \mathbb{E}[e^{-\eta X_T} I_{\{\tau=T\}}] &\leq \mathbb{E}[e^{-\eta R_0 e^{RT}} e^{-\eta \int_0^T e^{R(t-s)} p ds} e^{-\eta \int_0^T e^{R(t-s)} \sigma_0 dW_s}] \\ &\leq \mathbb{E}[e^{-\eta \int_0^T e^{R(t-s)} \sigma_0 dW_s}] \\ &= e^{\frac{\eta^2}{2} \int_0^T e^{2R(t-s)} \sigma_0^2 ds} < +\infty, \end{aligned}$$

and denoting by C a generic constant (possibly different from each line to another)

$$\begin{aligned} &\mathbb{E}[e^{-\eta(X_\tau - K)e^{R(T-\tau)}} e^{-\eta \int_\tau^T e^{R(T-s)}(p-q+qu_s)ds} e^{-\eta \int_\tau^T e^{R(T-s)}\sigma_0 u_s dW_s} I_{\{\tau < T\}}] \\ &\leq C \times \mathbb{E}[e^{-\eta(X_\tau - K)e^{R(T-\tau)}} e^{-\eta \int_\tau^T e^{R(T-s)}\sigma_0 u_s dW_s} I_{\{\tau < T\}}] \\ &\leq C \times \left(\mathbb{E}[e^{-2\eta(X_\tau - K)e^{2R(T-\tau)}} I_{\{\tau < T\}}] + \mathbb{E}[e^{-2\eta \int_\tau^T e^{R(T-s)}\sigma_0 u_s dW_s} I_{\{\tau < T\}}] \right) \\ &\leq C \times \left(\mathbb{E}[e^{-2\eta X_\tau e^{2R(T-\tau)}}] + \mathbb{E}[e^{-2\eta \int_\tau^T e^{R(T-s)}\sigma_0 u_s dW_s} I_{\{\tau < T\}}] \right) \\ &\leq C \times \left(\mathbb{E}[e^{-2\eta e^{2R(T-\tau)} \int_0^T e^{R(t-s)} \sigma_0 dW_s}] + \mathbb{E}[e^{2\eta^2 \int_\tau^T e^{2R(T-s)} \sigma_0^2 u_s^2 I_{\{\tau < T\}} ds}] \right) \\ &\leq C \times \left(e^{2\eta^2 e^{4RT} \int_0^T e^{2R(t-s)} \sigma_0^2 ds} + e^{2\eta^2 \int_0^T e^{2R(T-s)} \sigma_0^2 ds} \right) < +\infty. \end{aligned}$$

118 \square

Let us introduce the value function associated to our problem (2.10):

$$V(t, x) = \inf_{\alpha \in \mathcal{A}_t} \mathbb{E}[e^{-\eta X_T^{\alpha, t, x}}], \quad (t, x) \in [0, T] \times \mathbb{R}, \quad (2.11)$$

where $X^{\alpha,t,x} = \{X_s^{\alpha,t,x}\}_{s \in [t,T]}$ denotes the wealth given in equation (2.9) with initial condition $(t, x) \in [0, T] \times \mathbb{R}$, that is $X_t^{\alpha,t,x} = x$. We notice that

$$V(T, x) = e^{-\eta x} \quad \forall x \in \mathbb{R}, \quad (2.12)$$

119 because $X_T^{\alpha,T,x} = x \quad \forall \alpha \in \mathcal{A}$.

120 3. The pure reinsurance problem

In order to have a self-contained article, in this section we briefly investigate a pure reinsurance problem, which corresponds to the problem (2.10) with fixed starting time $t = 0$. Precisely, we deal with

$$\inf_{u \in \mathcal{U}} \mathbb{E}[e^{-\eta X_T^u}],$$

where \mathcal{U} denotes the class of admissible strategies $u = \{u_t\}_{t \in [0,T]}$, which are all the \mathbb{F} -predictable processes with values in $[0, 1]$. Let us denote by $\bar{V}(t, x)$ the value function associated to this problem, that is

$$\bar{V}(t, x) = \inf_{u \in \mathcal{U}_t} \mathbb{E}[e^{-\eta X_T^{u,t,x}}], \quad (t, x) \in [0, T] \times \mathbb{R}, \quad (3.1)$$

with \mathcal{U}_t denoting the restriction of \mathcal{U} to the time interval $[t, T]$ and $\{X_s^{u,t,x}\}_{s \in [t,T]}$ denotes the process satisfying equation (2.5) with initial data $(t, x) \in [0, T] \times \mathbb{R}$. It is well known that the value function (3.1) can be characterized as a classical solution to the associated Hamilton-Jacobi-Bellman (HJB) equation:

$$\begin{cases} \min_{u \in [0,1]} \mathcal{L}^u \bar{V}(t, x) = 0, & \forall (t, x) \in [0, T] \times \mathbb{R} \\ \bar{V}(T, x) = e^{-\eta x} & \forall x \in \mathbb{R}, \end{cases} \quad (3.2)$$

where, using equations (2.4) and (2.5), the generator of the Markov process X^u is given by

$$\mathcal{L}^u f(t, x) = \frac{\partial f}{\partial t}(t, x) + (Rx + p - q + qu) \frac{\partial f}{\partial x}(t, x) + \frac{1}{2} \sigma_0^2 u^2 \frac{\partial^2 f}{\partial x^2}(t, x), \quad f \in \mathcal{C}^{1,2}([0, T] \times \mathbb{R}),$$

with $\mathcal{C}^{1,2}([0, T] \times \mathbb{R})$ denoting the class of continuous functions, with continuous first order partial derivative with respect to the first (time) variable and continuous second order derivative with respect to the second (space) variable.

Under the ansatz $\bar{V}(t, x) = e^{-\eta x e^{R(T-t)}} \phi(t)$, the HJB equation reads as

$$\phi'(t) + \Psi(t) \phi(t) = 0, \quad \phi(T) = 1,$$

where

$$\Psi(t) = \min_{u \in [0,1]} \{-\eta e^{R(T-t)} (p - q + qu) + \frac{1}{2} \sigma_0^2 u^2 \eta^2 e^{2R(T-t)}\}.$$

Solving the minimization problem we find the unique minimizer:

$$u^*(t) = \frac{q}{\eta \sigma_0^2} e^{-R(T-t)} \vee 1, \quad t \in [0, T].$$

Under the additional condition

$$q < \eta \sigma_0^2, \quad (3.3)$$

$u^*(t)$ simplifies to

$$u^*(t) = \frac{q}{\eta \sigma_0^2} e^{-R(T-t)} \in (0, 1), \quad t \in [0, T]. \quad (3.4)$$

Using this expression we readily obtain that

$$\Psi(t) = \eta e^{R(T-t)}(q - p) - \frac{1}{2} \frac{q^2}{\sigma_0^2}. \quad (3.5)$$

By classical verification arguments, we can verify that the value function given in (3.1) takes this form:

$$\begin{aligned} \bar{V}(t, x) &= e^{-\eta x e^{R(T-t)}} e^{\int_t^T \Psi(s) ds} \\ &= e^{-\eta x e^{R(T-t)}} e^{\frac{\eta(q-p)}{R}(e^{R(T-t)} - 1)} e^{-\frac{1}{2} \frac{q^2}{\sigma_0^2}(T-t)}, \end{aligned} \quad (3.6)$$

121 and, under the condition (3.3), equation (3.4) provides an optimal reinsurance strategy.

Remark 1. Comparing the optimal strategy $u^*(s)$, $s \in [t, T]$, to the null reinsurance $u(s) = 1$, $s \in [t, T]$, by means of (3.1) we get that

$$\bar{V}(t, x) \leq \mathbb{E}[e^{-\eta X_T^{t,x}}] \quad \forall (t, x) \in [0, T] \times \mathbb{R}. \quad (3.7)$$

Moreover, by equation (2.2) we have that

$$\begin{aligned} g(t, x) &\doteq \mathbb{E}[e^{-\eta X_T^{t,x}}] = \mathbb{E}\left[e^{-\eta x e^{R(T-t)}} e^{-\eta \int_t^T e^{R(T-s)} p ds} e^{-\eta \int_t^T e^{R(T-s)} \sigma_0 dW_s}\right] \\ &= e^{-\eta x e^{R(T-t)}} e^{-\frac{\eta}{R}(e^{R(T-t)} - 1)p} e^{\frac{1}{4R}\eta^2 \sigma_0^2 (e^{2R(T-t)} - 1)}. \end{aligned} \quad (3.8)$$

Defining

$$h(t) \doteq \eta e^{R(T-t)} \left(\frac{1}{2} \eta e^{R(T-t)} \sigma_0^2 - p \right), \quad (3.9)$$

we can write

$$g(t, x) = e^{-\eta x e^{R(T-t)}} e^{\int_t^T h(s) ds}. \quad (3.10)$$

Hence, using (3.5) and (3.6), inequality (3.7) reads as

$$\begin{aligned} &\int_t^T [\Psi(s) - h(s)] ds \\ &= \frac{\eta q}{R}(e^{R(T-t)} - 1) - \frac{1}{2} \frac{q^2}{\sigma_0^2}(T-t) - \frac{1}{4R}\eta^2 \sigma_0^2 (e^{2R(T-t)} - 1) \leq 0 \quad \forall t \in [0, T]. \end{aligned} \quad (3.11)$$

122 4. Reduction to an optimal stopping problem

123 We can show that the mixed stochastic control problem (2.11) can be reduced to an
124 optimal stopping problem. Let us denote by $\mathcal{T}_{t,T}$ is the set of \mathbb{F} -stopping times τ such
125 that $t \leq \tau \leq T$.

Theorem 1. We have that

$$V(t, x) = \inf_{\tau \in \mathcal{T}_{t,T}} \mathbb{E}[\bar{V}(\tau, X_\tau^{t,x}) - K] I_{\{\tau < T\}} + e^{-\eta X_\tau^{t,x}} I_{\{\tau = T\}} \quad \forall t \in [0, T] \times \mathbb{R}, \quad (4.1)$$

126 where \bar{V} is given in (3.1) and $X^{t,x} = \{X_s^{t,x}\}_{s \in [t, T]}$ denotes the wealth process given in equation
127 (2.2), with initial data $(t, x) \in [0, T] \times \mathbb{R}$.

128 Moreover, let $\tau_{t,x}^* \in \mathcal{T}_{t,T}$ an optimal stopping time for problem (4.1). Then $\alpha^* =$
129 $(\tau_{t,x}^*, \{u_s^*\}_{s \in [\tau_{t,x}^*, T]})$, with $u^* = \{u_t^*\}_{t \in [0, T]}$ given in (3.4), is an optimal strategy for problem
130 (2.11), with the convention that on the event $\{\tau_{t,x}^* = T\}$ we take $\alpha^* = (T, 1)$.

Proof. We first prove the inequality

$$V(t, x) \geq \inf_{\tau \in \mathcal{T}_{t,T}} \mathbb{E}[\bar{V}(\tau, X_{\tau}^{t,x} - K)I_{\{\tau < T\}} + e^{-\eta X_{\tau}^{t,x}} I_{\{\tau = T\}}].$$

131 For any arbitrary strategy $\alpha = (\tau, \{u_s\}_{s \in [\tau, T]}) \in \mathcal{A}_t$, we have that $\tau \in \mathcal{T}_{t,T}$ and by (2.9)

$$\mathbb{E}[e^{-\eta X_T^{\alpha, t, x}}] = \mathbb{E}[e^{-\eta X_T^{u, \tau, X_{\tau}^{t,x} - K}} I_{\{\tau < T\}} + e^{-\eta X_{\tau}^{t,x}} I_{\{\tau = T\}}] =$$

$$\mathbb{E}[\mathbb{E}[e^{-\eta X_T^{u, \tau, X_{\tau}^{t,x} - K}} | \mathcal{F}_{\tau}] I_{\{\tau < T\}} + e^{-\eta X_{\tau}^{t,x}} I_{\{\tau = T\}}] \geq$$

$$\mathbb{E}[\bar{V}(\tau, X_{\tau}^{t,x} - K)I_{\{\tau < T\}} + e^{-\eta X_{\tau}^{t,x}} I_{\{\tau = T\}}].$$

Taking the infimum over $\alpha \in \mathcal{A}_t$ on both side leads to the desired inequality. The other side of the inequality is based on the fact that there exists $u^* = \{u_t^*\}_{t \in [0, T]} \in \mathcal{U}$, given in (3.4) optimal for the problem (3.1). Indeed, consider the strategy $\bar{\alpha} = (\tau, \{u_s^*\}_{s \in [\tau, T]}) \in \mathcal{A}_t$ where τ is arbitrary chosen in $\mathcal{T}_{t,T}$. Then

$$V(t, x) \leq \mathbb{E}[e^{-\eta X_T^{\bar{\alpha}, t, x}}] = \mathbb{E}[e^{-\eta X_T^{u^*, \tau, X_{\tau}^{t,x} - K}} I_{\{\tau < T\}} + e^{-\eta X_{\tau}^{t,x}} I_{\{\tau = T\}}] =$$

$$\mathbb{E}[\bar{V}(\tau, X_{\tau}^{t,x} - K)I_{\{\tau < T\}} + e^{-\eta X_{\tau}^{t,x}} I_{\{\tau = T\}}].$$

Taking the infimum over $\tau \in \mathcal{T}_{t,T}$ on the right-hand side gives that

$$V(t, x) \leq \inf_{\tau \in \mathcal{T}_{t,T}} \mathbb{E}[\bar{V}(\tau, X_{\tau}^{t,x} - K)I_{\{\tau < T\}} + e^{-\eta X_{\tau}^{t,x}} I_{\{\tau = T\}}] \quad \forall t \in [0, T] \times \mathbb{R},$$

132 and hence the equality (4.1).

Finally, let $\tau_{t,x}^* \in \mathcal{T}_{t,T}$ an optimal stopping time for problem (4.1) and $a^* = (\tau_{t,x}^*, \{u_s^*\}_{s \in [\tau_{t,x}^*, T]})$ with $u^* = \{u_t^*\}_{t \in [0, T]}$ given in (3.4), we get that

$$V(t, x) = \mathbb{E}[\bar{V}(\tau_{t,x}^*, X_{\tau_{t,x}^*}^{t,x} - K)I_{\{\tau_{t,x}^* < T\}} + e^{-\eta X_{\tau_{t,x}^*}^{t,x}} I_{\{\tau_{t,x}^* = T\}}]$$

$$= \mathbb{E}[e^{-\eta X_T^{u^*, \tau_{t,x}^*, X_{\tau_{t,x}^*}^{t,x} - K}} I_{\{\tau_{t,x}^* < T\}} + e^{-\eta X_{\tau_{t,x}^*}^{t,x}} I_{\{\tau_{t,x}^* = T\}}]$$

$$= \mathbb{E}[e^{-\eta X_T^{a^*, t, x}}],$$

133 and this concludes the proof. \square

134 According to Theorem 1 we can solve the original problem given in (2.11) in two
135 steps: after investigating the pure reinsurance problem (3.1) (see Section 3), we can
136 analyze the optimal stopping problem (4.1), which is the main goal of the next section.

137 5. The optimal stopping problem

In this section we discuss the optimal stopping problem (4.1):

$$V(t, x) = \inf_{\tau \in \mathcal{T}_{t,T}} \mathbb{E}[\bar{V}(\tau, X_{\tau}^{t,x} - K)I_{\{\tau < T\}} + e^{-\eta X_{\tau}^{t,x}} I_{\{\tau = T\}}] \quad t \in [0, T] \times \mathbb{R}.$$

Let us observe that

$$V(T, x) = e^{-\eta x} \quad \forall x \in \mathbb{R},$$

while choosing $\tau = t < T$ and $\tau = T$ in the right hand side of (4.1), we get that

$$V(t, x) \leq \bar{V}(t, x - K) \quad \text{and} \quad V(t, x) \leq \mathbb{E}[e^{-\eta X_T^{t,x}}] = g(t, x), \forall (t, x) \in [0, T] \times \mathbb{R}, \quad (5.1)$$

138 respectively.

139

Now denote by \mathcal{L} the Markov generator of the process $X^{t,x}$:

$$\mathcal{L}f(t, x) = \frac{\partial f}{\partial t}(t, x) + (Rx + p) \frac{\partial f}{\partial x}(t, x) + \frac{1}{2} \sigma_0^2 \frac{\partial^2 f}{\partial x^2}(t, x), \quad (5.2)$$

140 with $f \in \mathcal{C}^{1,2}((0, T) \times \mathbb{R})$.

Remark 2. From the theory of optimal stopping (see, for instance [17]), when the cost function $G(t, x)$ is continuous and the value function

$$W(t, x) = \inf_{\tau \in \mathcal{T}_{t,T}} \mathbb{E}[G(\tau, X_\tau^{t,x})], \quad t \in [0, T] \times \mathbb{R}$$

is sufficiently regular, it can be characterized as a solution to the following variational inequality:

$$\min\{\mathcal{L}W(t, x), G(t, x) - W(t, x)\} = 0, \quad (t, x) \in (0, T) \times \mathbb{R}. \quad (5.3)$$

This is a free-boundary problem, whose solution is the function $W(t, x)$ and the so-called continuation region, which is defined as

$$\mathcal{C} = \{(t, x) \in (0, T) \times \mathbb{R} : W(t, x) < G(t, x)\}. \quad (5.4)$$

Moreover, it is known that the first exit time of the process $X^{t,x}$ from the region \mathcal{C}

$$\tau_{t,x}^* \doteq \inf\{s \in [t, T] : (s, X_s^{t,x}) \notin \mathcal{C}\}.$$

141 provides an optimal stopping time.

In our optimal stopping problem (4.1), the cost function is

$$\bar{V}(t, x - K) I_{\{t < T\}} + e^{-\eta x} I_{\{t = T\}}, \quad t \in [0, T] \times \mathbb{R},$$

142 which is not continuous on $[0, T] \times \mathbb{R}$, hence the classical theory on optimal stopping problems
143 does not directly apply.

144 In view of the preceding remark, we now prove a Verification Theorem which
145 applies to our specific problem.

Theorem 2 (Verification Theorem). Let $\varphi: [0, T] \times \mathbb{R} \rightarrow \mathbb{R}$ be a function satisfying the assumptions below and \mathcal{C} (the continuation region) be defined by

$$\mathcal{C} = \{(t, x) \in (0, T) \times \mathbb{R} : \varphi(t, x) < \bar{V}(t, x - K)\}. \quad (5.5)$$

146 Suppose that the following conditions are satisfied.

- 147 1. There exists $t^* \in [0, T)$ such that $\mathcal{C} = (t^*, T) \times \mathbb{R}$.
- 148 2. $\varphi \in \mathcal{C}([0, T] \times \mathbb{R})$, φ is \mathcal{C}^1 w.r.t t in $(0, t^*)$ and (t^*, T) , separately, and \mathcal{C}^2 w.r.t. $x \in \mathbb{R}$;
- 149 3. $\varphi(t, x) \leq \bar{V}(t, x - K) \forall (t, x) \in [0, T] \times \mathbb{R}$ and $\varphi(T, x) = e^{-\eta x} \forall x \in \mathbb{R}$;
4. φ is a solution to the following variational inequality

$$\begin{cases} \mathcal{L}\varphi(t, x) \geq 0 & \forall (t, x) \in (0, t^*) \times \mathbb{R} \\ \mathcal{L}\varphi(t, x) = 0 & \forall (t, x) \in \mathcal{C} = (t^*, T) \times \mathbb{R}. \end{cases} \quad (5.6)$$

150 5. the family $\{\varphi(\tau, X_\tau); \tau \in \mathcal{T}_{0,T}\}$ is uniformly integrable.

Moreover, let $\tau_{t,x}^*$ the first exit time from the region \mathcal{C} of the process $X^{t,x}$, that is

$$\tau_{t,x}^* \doteq \inf\{s \in [t, T] : (s, X_s^{t,x}) \notin \mathcal{C}\}.$$

151 with the convention $\tau_{t,x}^* = T$ if the set on the right-hand side is empty.

152 Then $\varphi(t, x) = V(t, x)$ on $[0, T] \times \mathbb{R}$ and $\tau_{t,x}^*$ is an optimal stopping time for problem (4.1).

153 **Proof.** For any $(t, x) \in [0, T] \times \mathbb{R}$ let us take the sequence of stopping times $\{\tau_n\}_{n \geq 1}$
 154 such that $\tau_n \doteq \inf\{s \geq t \mid |X_s^{t,x}| \geq n\}$. We first prove that, $\forall \tau \in \mathcal{T}_{t,T}$

$$\varphi(t, x) \leq \mathbb{E}[\varphi(\tau \wedge \tau_n, X_{\tau \wedge \tau_n}^{t,x})], \quad \forall (t, x) \in [0, T] \times \mathbb{R}. \quad (5.7)$$

Due to the specific form of the continuation region we have two cases. If $t \geq t^*$, since $\varphi \in \mathcal{C}^{1,2}((t^*, T) \times \mathbb{R})$, applying Dynkin's formula¹ we get that for any arbitrary stopping time $\tau \in \mathcal{T}_{t,T}$

$$\varphi(t, x) = \mathbb{E}[\varphi(\tau \wedge \tau_n, X_{\tau \wedge \tau_n}^{t,x})] - \mathbb{E}\left[\int_t^{\tau \wedge \tau_n} \mathcal{L}\varphi(s, X_s^{t,x}) ds\right] = \mathbb{E}[\varphi(\tau \wedge \tau_n, X_{\tau \wedge \tau_n}^{t,x})].$$

If $t < t^*$, we have again by Dynkin's formula, since $\varphi \in \mathcal{C}^{1,2}((0, t^*) \times \mathbb{R})$, that

$$\begin{aligned} \varphi(t, x) &= \mathbb{E}[\varphi(\tau \wedge \tau_n \wedge t^*, X_{\tau \wedge \tau_n \wedge t^*}^{t,x})] - \mathbb{E}\left[\int_t^{\tau \wedge \tau_n \wedge t^*} \mathcal{L}\varphi(s, X_s^{t,x}) ds\right] \\ &\leq \mathbb{E}[\varphi(\tau \wedge \tau_n \wedge t^*, X_{\tau \wedge \tau_n \wedge t^*}^{t,x})] \end{aligned}$$

and, similarly, since $\varphi \in \mathcal{C}^{1,2}((t^*, T) \times \mathbb{R})$,

$$\begin{aligned} \mathbb{E}[\varphi(\tau \wedge \tau_n \wedge t^*, X_{\tau \wedge \tau_n \wedge t^*}^{t,x})] &= \mathbb{E}[\varphi(\tau \wedge \tau_n, X_{\tau \wedge \tau_n}^{t,x})] - \mathbb{E}\left[\int_{\tau \wedge \tau_n \wedge t^*}^{\tau \wedge \tau_n} \mathcal{L}\varphi(s, X_s^{t,x}) ds\right] \\ &= \mathbb{E}[\varphi(\tau \wedge \tau_n, X_{\tau \wedge \tau_n}^{t,x})], \end{aligned}$$

155 hence (5.7) is proved.

Now letting $n \rightarrow +\infty$ in (5.7), recalling that $\varphi \in \mathcal{C}([0, T] \times \mathbb{R})$ and using Fatou's Lemma we get that

$$\varphi(t, x) \leq \mathbb{E}[\varphi(\tau, X_\tau^{t,x})] \leq \mathbb{E}[\bar{V}(\tau, X_\tau^{t,x} - K)I_{\{\tau < T\}} + e^{-\eta X_\tau^{t,x}} I_{\{\tau = T\}}] \quad \forall \tau \in \mathcal{T}_{t,T},$$

156 hence $\varphi(t, x) \leq V(t, x)$, $\forall (t, x) \in [0, T] \times \mathbb{R}$. To prove the opposite inequality we
 157 consider four different cases.

- 158 1. If the stopping region is not empty, that is $t^* \in (0, T)$, $\forall (t, x) \in (0, t^*) \times \mathbb{R}$ we
 159 know that $\varphi(t, x) = \bar{V}(t, x - K) \geq V(t, x)$, hence $\varphi(t, x) = V(t, x)$, which implies
 160 $V(t, x) = \bar{V}(t, x - K)$ and $\tau_{t,x}^* = t$ is optimal for problem (4.1).
- 161 2. If the stopping region is not empty, for $t = 0$, we have that $\varphi(0, x) = \bar{V}(0, x - K)$
 162 $\forall x \in \mathbb{R}$, otherwise by continuity of both the functions if $\varphi(0, x) > \bar{V}(0, x - K)$ (or
 163 $\varphi(0, x) < \bar{V}(0, x - K)$) the same inequality holds in a neighborhood of $(0, x)$ which
 164 contradicts that $\varphi(t, x) = \bar{V}(t, x - K)$, $\forall (t, x) \in (0, t^*) \times \mathbb{R}$. Then $\varphi(0, x) = V(0, x)$
 165 $\forall x \in \mathbb{R}$ and $\tau_{0,x}^* = 0$ is optimal for problem (4.1).

¹ Notice that we use a localization argument, so that $\tau \wedge \tau_n$ is the first exit time of a bounded set and, as a consequence, φ is not required to have a compact support (see [17, Theorem 7.4.1]).

3. If the continuation region is not empty, that is $t^* \in [0, T)$, $\forall (t, x) \in [t^*, T) \times \mathbb{R}$, repeating the localization argument with the stopping time $\tau_{t,x}^* = T$, we get

$$\begin{aligned}\varphi(t, x) &= \mathbb{E}[\varphi(T, X_T^{t,x})] - \mathbb{E}\left[\int_t^T \mathcal{L}\varphi(s, X_s^{t,x}) ds\right] \\ &= \mathbb{E}[\varphi(T, X_T^{t,x})] = \mathbb{E}[e^{-\eta X_T^{t,x}}] \geq V(t, x),\end{aligned}$$

166 as a consequence $\varphi(t, x) = V(t, x) = \mathbb{E}[e^{-\eta X_T^{t,x}}]$ and $\tau_{t,x}^* = T$ is optimal for problem
167 (4.1).

- 168 4. Finally, for $t = T$ by assumption $\varphi(T, x) = e^{-\eta x} = V(T, x)$, $\forall x \in \mathbb{R}$, $\tau_{T,x}^* = T$ is
169 optimal for problem (4.1) and this concludes the proof.

170 \square

171 **Lemma 1.** Let g as defined in equation (3.10). The families $\{\bar{V}(\tau, X_\tau - K); \tau \in \mathcal{T}_{0,T}\}$ and
172 $\{g(\tau, X_\tau); \tau \in \mathcal{T}_{0,T}\}$ are uniformly integrable.

Proof. Recalling that $\bar{V}(t, x) \leq g(t, x)$ by (3.7), we have that $\bar{V}(t, x - k) \leq e^{\eta K e^{RT}} g(t, x)$, hence the statement follows by the uniform integrability of the family

$$\{g(\tau, X_\tau) : \tau \in \mathcal{T}_{0,T}\}.$$

It is well known that if for any arbitrary $\delta > 0$ and any stopping time $\tau \in \mathcal{T}_{0,T}$

$$\mathbb{E}[g(\tau, X_\tau)^{1+\delta}] < +\infty,$$

then the proof is complete. To this end, we observe that

$$\begin{aligned}\mathbb{E}[g(\tau, X_\tau)^{1+\delta}] &= \mathbb{E}[e^{(1+\delta) \int_\tau^T h(s) ds} e^{-\eta(1+\delta)e^{R(T-\tau)} X_\tau}] \\ &\leq e^{\frac{1}{4R}(1+\delta)\eta^2\sigma_0^2 e^{2RT}} e^{-\eta(1+\delta)R_0 e^{RT}} \mathbb{E}[e^{-\eta(1+\delta) \int_0^\tau e^{R(T-s)} \sigma_0 dW_s}] \\ &\leq e^{\frac{1}{4R}(1+\delta)\eta^2\sigma_0^2 e^{2RT}} e^{-\eta(1+\delta)R_0 e^{RT}} e^{\frac{1}{4R}(1+\delta)^2\eta^2\sigma_0^2 (e^{2RT} - 1)} < +\infty.\end{aligned}$$

173 \square

174 The guess for the continuation region \mathcal{C} given in the assumption 1. of the Verification
175 Theorem follows by the next result.

Lemma 2. The set

$$A \doteq \{(t, x) \in (0, T) \times \mathbb{R} : \mathcal{L}\bar{V}(t, x - K) < 0\} \quad (5.8)$$

is included in the continuation region, that is

$$A \subseteq \mathcal{C} = \{(t, x) \in (0, T) \times \mathbb{R} : V(t, x) < \bar{V}(t, x - K)\}.$$

Moreover, the following equation holds:

$$A = (t_A, T) \times \mathbb{R},$$

where

$$t_A \doteq 0 \vee \left[T - \frac{1}{R} \log \left(\frac{q + RK + \sqrt{(q + RK)^2 - q^2}}{\eta\sigma_0^2} \right) \right] \wedge T. \quad (5.9)$$

176 In particular, only three cases are possible, depending on the model parameters:

1. if

$$y^* := \frac{q + RK + \sqrt{(q + RK)^2 - q^2}}{\eta\sigma_0^2} \geq e^{RT},$$

177 then $t_A = 0$ and $\mathcal{L}\bar{V}(t, x - K) < 0 \forall (t, x) \in (0, T) \times \mathbb{R}$, so that $A = (0, T) \times \mathbb{R}$,
178 implying that $\mathcal{C} = (0, T) \times \mathbb{R}$;

2. if

$$1 < y^* = \frac{q + RK + \sqrt{(q + RK)^2 - q^2}}{\eta\sigma_0^2} < e^{RT},$$

179 then $0 < t_A < T$ and $\mathcal{L}\bar{V}(t, x - K) < 0 \forall (t, x) \in (t_A, T) \times \mathbb{R}$; in this case $A =$
180 $(t_A, T) \times \mathbb{R}$;

3. if

$$y^* = \frac{q + RK + \sqrt{(q + RK)^2 - q^2}}{\eta\sigma_0^2} \leq 1,$$

181 then $t_A = T$ and $\mathcal{L}\bar{V}(t, x - K) \geq 0 \forall (t, x) \in (0, T) \times \mathbb{R}$, so that $A = \emptyset$.

Proof. First let us observe that $\bar{V}(t, x - K) \in C^{1,2}((0, T) \times \mathbb{R}) \cap \mathcal{C}([0, T] \times \mathbb{R})$ and the family $\{\bar{V}(\tau, X_\tau - K); \tau \in \mathcal{T}_{0,T}\}$ is uniformly integrable by Lemma 1. Now choose $(\bar{t}, \bar{x}) \in A$, let $B \subset A$ be a neighborhood of (\bar{t}, \bar{x}) with $\tau_B < T$, where τ_B denotes the first exit time of $X^{\bar{t}, \bar{x}}$ from B . Then by Dynkin's formula

$$\begin{aligned} \bar{V}(\bar{t}, \bar{x} - K) &= \mathbb{E}[\bar{V}(\tau_B, X_{\tau_B}^{\bar{t}, \bar{x}} - K)] - \mathbb{E}\left[\int_{\bar{t}}^{\tau_B} \mathcal{L}\bar{V}(s, X_s^{\bar{t}, \bar{x}} - K) ds\right] \\ &> \mathbb{E}[\bar{V}(\tau_B, X_{\tau_B}^{\bar{t}, \bar{x}} - K)] \geq V(\bar{t}, \bar{x}). \end{aligned}$$

182 Hence $(\bar{t}, \bar{x}) \in \mathcal{C}$ and $A \subseteq \mathcal{C}$.

Next, recalling (5.2), we have that

$$\mathcal{L}\bar{V}(t, x - K) = \bar{V}(t, x - K)(-\Psi(t) - \eta e^{R(T-t)}(p + RK) + \frac{1}{2}\eta^2 e^{2R(T-t)}\sigma_0^2),$$

so that $\mathcal{L}\bar{V}(t, x - K) < 0$ if and only if

$$\Psi(t) > \frac{1}{2}\eta^2 e^{2R(T-t)}\sigma_0^2 - \eta e^{R(T-t)}(p + RK), \quad (5.10)$$

that is, using (3.5),

$$\frac{1}{2}\eta^2 e^{2R(T-t)}\sigma_0^2 - \eta e^{R(T-t)}(q + RK) + \frac{1}{2}\frac{q^2}{\sigma_0^2} < 0.$$

Using a change of variable $z = e^{R(T-t)}$, we can rewrite the inequality as

$$\frac{1}{2}\eta^2\sigma_0^2 z^2 - \eta(q + RK)z + \frac{1}{2}\frac{q^2}{\sigma_0^2} < 0.$$

Since $\eta^2[(q + K)^2 - q^2] > 0$ the associated equation admits two different solutions, so that the inequality (5.10) is satisfied by

$$\frac{q + RK - \sqrt{(q + RK)^2 - q^2}}{\eta\sigma_0^2} < z < \frac{q + RK + \sqrt{(q + RK)^2 - q^2}}{\eta\sigma_0^2}.$$

Recalling (3.3), we can verify that

$$\frac{q + RK - \sqrt{(q + RK)^2 - q^2}}{\eta\sigma_0^2} < \frac{q}{\eta\sigma_0^2} < 1,$$

so that the inequality reads as

$$t_A = T - \frac{1}{R} \log \left(\frac{q + RK + \sqrt{(q + RK)^2 - q^2}}{\eta\sigma_0^2} \right) < t < T.$$

183 Depending on the model parameters, we can see that only the three cases above are
184 possible. Equivalently, $\mathcal{L}\tilde{V}(t, x - K) < 0$ if and only if $t_A < t < T$. \square

Remark 3. As consequence of Lemma 2, recalling (3.9), in Cases 1 and 2, that is when $0 \leq t_A < T$, we have that

$$\Psi(t) - h(t) + \eta R K e^{R(T-t)} > 0, \quad \forall t > t_A,$$

see equation (5.10), which implies, $\forall t \geq t_A$

$$\int_t^T (\Psi(s) - h(s)) ds + \int_t^T \eta R K e^{R(T-s)} ds > 0,$$

equivalently

$$\int_t^T (\Psi(s) - h(s)) ds + \eta K e^{R(T-t)} > \eta K$$

for all $t \in [t_A, T)$. In Case 3, that is when $t_A = T$, since

$$\Psi(t) - h(t) + \eta R K e^{R(T-t)} < 0, \quad \forall t \in [0, T)$$

we have that

$$\int_t^T (\Psi(s) - h(s)) ds + \eta K e^{R(T-t)} < \eta K,$$

185 for all $t \in [0, T)$.

186 We need the following preliminary result to provide an explicit expression for the
187 value function of the problem (4.1).

188 **Lemma 3.** The function $\tilde{V}(t, x) = Cg(t, x)$, $(t, x) \in (0, T) \times \mathbb{R}$, with C any positive con-
189 stant and g as given in equation (3.10), is a solution to the partial differential equation (PDE)
190 $\mathcal{L}\tilde{V}(t, x) = 0$, $(t, x) \in (0, T) \times \mathbb{R}$.

191 In particular, g is a solution to the PDE with boundary condition $g(T, x) = e^{-\eta x} \forall x \in \mathbb{R}$.

Proof. Using the ansatz $\tilde{V}(t, x) = e^{-\eta x e^{R(T-t)}} \gamma(t)$, we can reduce the PDE $\mathcal{L}\tilde{V}(t, x) = 0$ to the following equation:

$$e^{-\eta x e^{R(T-t)}} \gamma'(t) - \eta e^{R(T-t)} V(t, x) p + \frac{1}{2} \eta^2 e^{2R(T-t)} V(t, x) \sigma_0^2 = 0,$$

which is equivalent to this ordinary differential equation (ODE):

$$\gamma'(t) + h(t)\gamma(t) = 0, \quad (t, x) \in (0, T) \times \mathbb{R},$$

192 where the function h is given in (3.9).

193 Since the solution of the ODE is $\gamma(t) = C e^{\int_t^T h(s) ds}$, we get the expression of \tilde{V} as
194 above.

195 Finally, setting $C = 1$, g satisfies the PDE above with the terminal condition $g(T, x) =$
 196 $e^{-\eta x} \forall x \in \mathbb{R}$. \square

197 Before proving the main result of this section, which is Theorem 3, we compare
 198 $g(t, x)$, given in (3.10), with $\bar{V}(t, x - K)$.

Lemma 4. *Let*

$$H(t) = \int_t^T (\Psi(s) - h(s)) ds + \eta K e^{R(T-t)}, \quad t \in [0, T], \quad (5.11)$$

199 then we distinguish two cases:

- 200 1. if $H(0) \geq 0$, then $g(t, x) < \bar{V}(t, x - K) \forall (t, x) \in (0, T] \times \mathbb{R}$;
 201 2. if $H(0) < 0$, then there exists $t^* \in (0, t_A)$ such that $g(t^*, x) = \bar{V}(t^*, x - K) \forall x \in \mathbb{R}$
 202 and $g(t, x) < \bar{V}(t, x - K) \forall (t, x) \in (t^*, T] \times \mathbb{R}$.

Proof. Let us observe that the inequality $g(t, x) < \bar{V}(t, x - K)$ writes as

$$e^{-\eta x e^{R(T-t)}} e^{\int_t^T h(s) ds} < e^{-\eta(x-K)e^{R(T-t)}} e^{\int_t^T \Psi(s) ds},$$

that is

$$e^{\int_t^T (\Psi(s) - h(s)) ds} e^{\eta K e^{R(T-t)}} > 1 \Leftrightarrow \int_t^T (\Psi(s) - h(s)) ds + \eta K e^{R(T-t)} = H(t) > 0.$$

203 We distinguish three cases:

- 204 (i) when $0 \leq t_A < T$, we have that $H(t) \geq \eta K > 0 \forall t > t_A$ by Remark 3 and it easy
 205 to verify that H is increasing in $[0, t_A]$, while it is decreasing in $[t_A, T]$. Hence, it
 206 takes the maximum value at $t = t_A$. As a consequence, if $H(0) \geq 0$ we have that
 207 $H(t) > 0 \forall t \in (0, T]$, being $H(T) = \eta K > 0$.
 208 Otherwise, if $H(0) < 0$ there exists $t^* \in (0, t_A)$ such that $H(t^*) = 0$, that is
 209 $g(t^*, x) = \bar{V}(t^*, x - K) \forall x \in \mathbb{R}$, and $H(t) > 0 \forall (t, x) \in (t^*, T]$, that is $g(t, x) <$
 210 $\bar{V}(t, x - K) \forall (t, x) \in (t^*, T] \times \mathbb{R}$;
 211 (ii) when $t_A = T$, by Lemma 2 we get that H is increasing in $[0, T]$ and we can repeat
 212 the same arguments as in the previous case to distinguish the two cases $H(0) \geq 0$
 213 and $H(0) < 0$, obtaining the same results;
 214 (iii) when $t_A = 0$, by Remark 3 we know that H is decreasing in $[0, T]$, so that $H(t) \geq$
 215 $\eta K > 0 \forall t \in [0, T]$, that is $g(t, x) < \bar{V}(t, x - K), \forall (t, x) \in (0, T] \times \mathbb{R}$. Moreover, in
 216 this case $H(0) \geq 0$.

217 Summarizing, we obtain our statement. \square

218 We now prove some properties of the continuation region.

Proposition 2. *Let*

$$\mathcal{C} = \{(t, x) \in (0, T) \times \mathbb{R} : V(t, x) < \bar{V}(t, x - K)\}. \quad (5.12)$$

219 Then we distinguish two cases:

- 220 1. if $H(0) \geq 0$, then $\mathcal{C} = (0, T) \times \mathbb{R}$,
 2. if $H(0) < 0$, then $(t^*, T) \times \mathbb{R} \subseteq \mathcal{C}$, where $t^* \in (0, t_A)$ is the unique solution to equation

$$H(t) = \int_t^T (\Psi(s) - h(s)) ds + \eta K e^{R(T-t)} = 0.$$

221 **Proof.** We apply Lemma 4. In Case 1, we have that $V(t, x) \leq g(t, x) < \bar{V}(t, x - K)$
 222 $\forall (t, x) \in (0, T) \times \mathbb{R}$, that is $\mathcal{C} = (0, T) \times \mathbb{R}$. In Case 2, we have that $V(t, x) \leq g(t, x) <$

223 $\bar{V}(t, x - K) \forall (t, x) \in (t^*, T) \times \mathbb{R}$, which implies $(t^*, T) \times \mathbb{R} \subseteq \mathcal{C}$, and this concludes the
224 proof. \square

225 Now we are ready for the main result of this section.

226 **Theorem 3.** Let H be given in (5.11). The solution of the optimal stopping problem (4.1) takes
227 different forms, depending on the model parameters. Precisely, we have two cases:

1. if $H(0) = \int_0^T (\psi(s) - h(s))ds + \eta Ke^{RT} \geq 0$, then the continuation region is $\mathcal{C} = (0, T) \times \mathbb{R}$, the value function is

$$V(t, x) = g(t, x) = e^{-\eta xe^{R(T-t)}} e^{\int_t^T h(s) ds}, \quad (t, x) \in [0, T] \times \mathbb{R}$$

228 and $\tau_{t,x}^* = T$ is an optimal stopping time;

2. if $H(0) = \int_0^T (\psi(s) - h(s))ds + \eta Ke^{RT} < 0$, then $\mathcal{C} = (t^*, T) \times \mathbb{R}$, where $t^* \in (t_A, T)$ is the unique solution to $H(t) = 0$, the value function is

$$V(t, x) = \begin{cases} \bar{V}(t, x - K) = e^{-\eta(x-K)e^{R(T-t)}} e^{\int_t^T \Psi(s) ds} & (t, x) \in [0, t^*] \times \mathbb{R} \\ g(t, x) = \mathbb{E}[e^{-\eta X_T^{t,x}}] = e^{-\eta xe^{R(T-t)}} e^{\int_t^T h(s) ds} & (t, x) \in (t^*, T) \times \mathbb{R} \end{cases} \quad (5.13)$$

and $\tau_{t,x}^*$, given by

$$\tau_{t,x}^* = \begin{cases} t & (t, x) \in [0, t^*] \times \mathbb{R} \\ T & (t, x) \in (t^*, T) \times \mathbb{R}, \end{cases} \quad (5.14)$$

229 is an optimal stopping time.

230 **Proof.** We prove the two cases separately, applying Theorem 2 in each one.

231

Case 1

The continuation region is $\mathcal{C} = (0, T) \times \mathbb{R}$ by Proposition 2, hence assumption 1 of Theorem 2 is fulfilled. Moreover, $\tau_{t,x}^* = T$. Observing that

$$g(t, x) = e^{-\eta xe^{R(T-t)}} e^{\int_t^T h(s) ds} \in \mathcal{C}^{1,2}((0, T) \times \mathbb{R}) \cap \mathcal{C}([0, T] \times \mathbb{R}),$$

232 the assumption 2 of Theorem 2 is clearly matched. The assumption 3 is implied by
233 Lemma 4. Moreover, the variational inequality (5.6) (assumption 4) is fulfilled by Lemma
234 3. Finally, by Lemma 1 the last condition in Theorem 2 is fulfilled.

Case 2

$\mathcal{C} = (t^*, T) \times \mathbb{R}$ clearly satisfies the first assumption of Theorem 2. Taking

$$\varphi(t, x) = \begin{cases} \bar{V}(t, x - K) = e^{-\eta(x-K)e^{R(T-t)}} e^{\int_t^T \Psi(s) ds} & (t, x) \in [0, t^*] \times \mathbb{R} \\ g(t, x) = \mathbb{E}[e^{-\eta X_T^{t,x}}] = e^{-\eta xe^{R(T-t)}} e^{\int_t^T h(s) ds} & (t, x) \in (t^*, T) \times \mathbb{R}, \end{cases}$$

235 observing that Lemma 4 ensures the existence of $t^* \in (0, t_A)$ such that $g(t^*, x) =$
236 $\bar{V}(t^*, x - K)$ when $H(0) < 0$, the smoothness conditions of the second assumption are
237 matched. Moreover, according to Lemma 4, $g(t, x) < \bar{V}(t, x - K) \forall (t, x) \in (t^*, T]$ and
238 the assumption 3 is fulfilled. That the variational inequality (5.6) is satisfied by φ is a
239 consequence of the results of Section 3 and of Lemma 3. Finally, Lemma 1 implies the
240 fifth assumption of Theorem 2 and the proof is complete. \square

241 6. Solution to the original problem

242 As a direct consequence of the results obtained in the previous section and Theorem
243 1, we provide an explicit solution to the optimal reinsurance problem under fixed cost
244 given in (2.11).

Theorem 4. *Let us define*

$$K^* = -\frac{q}{R}(1 - e^{-RT}) + \frac{1}{2} \frac{q^2}{\eta\sigma_0^2} T e^{-RT} + \frac{1}{4R} \eta\sigma_0^2 (e^{RT} - e^{-RT}) > 0. \quad (6.1)$$

245 *Two cases are possible, depending on the model parameters:*

1. *if $K \geq K^*$, then the value function given in (2.11) is*

$$V(t, x) = g(t, x) = \mathbb{E}[e^{-\eta X_T^{t,x}}]$$

246 *and the optimal strategy is $\alpha^* = (T, 1)$, that is no reinsurance is purchased;*

2. *if $K < K^*$, then the value function is*

$$V(t, x) = \begin{cases} \bar{V}(t, x - K) = e^{-\eta(x-K)e^{R(T-t)}} e^{\int_t^T \Psi(s) ds} & (t, x) \in [0, t^*] \times \mathbb{R} \\ g(t, x) = \mathbb{E}[e^{-\eta X_T^{t,x}}] = e^{-\eta x e^{R(T-t)}} e^{\int_t^T h(s) ds} & (t, x) \in (t^*, T] \times \mathbb{R}, \end{cases}$$

where $t^* \in (0, T)$ is the unique solution to the equation

$$\eta \left(\frac{q}{R} + K \right) e^{R(T-t)} - \frac{1}{2} \frac{q^2}{\sigma_0^2} (T - t) - \frac{1}{4R} \eta^2 \sigma_0^2 (e^{2R(T-t)} - 1) - \frac{\eta q}{R} = 0,$$

247 *and the optimal strategy is $\alpha^* = (\tau_{t,x}^*, \{ \frac{q}{\eta\sigma_0^2} e^{-R(T-s)} \}_{s \in [\tau_{t,x}^*, T]})$, with $\tau_{t,x}^*$ given in (5.14).*

Proof. Let us observe that, using Remark 1,

$$\begin{aligned} H(t) &= \int_t^T (\Psi(s) - h(s)) ds + \eta K e^{R(T-t)} \\ &= \frac{\eta q}{R} (e^{R(T-t)} - 1) - \frac{1}{2} \frac{q^2}{\sigma_0^2} (T - t) - \frac{1}{4R} \eta^2 \sigma_0^2 (e^{2R(T-t)} - 1) + \eta K e^{R(T-t)}. \end{aligned}$$

and the condition $H(0) \geq 0$ is equivalent to

$$K \geq -\frac{1}{\eta} e^{-RT} \int_0^T (\Psi(s) - h(s)) ds = K^*,$$

248 while the condition $H(0) < 0$ can be written as $K < K^*$. That $K^* > 0$ follows by Remark

249 1. Then the statement is a consequence of Theorem 3. \square

250 Let us briefly comment the two cases of Theorem 4. Case 1 corresponds to no
251 reinsurance. That is, the insurer is not willing to subscribe a contract at any time of
252 the selected time horizon. Besides the insurer, this result is relevant for the reinsurance
253 company. We have proven that there exists a threshold $K^* > 0$ (see equation (6.1)), which
254 represents the maximum initial cost that the insurer is willing to pay to buy reinsurance.
255 If the reinsurer chooses a subscription cost higher than K^* , then the insurer will not buy
256 protection from her.

257 In Case 2, at any time $t \in [0, T]$, the insurer immediately subscribes the reinsurance
258 agreement if the time instant t^* has not passed, applying the optimal retention level
259 from that moment on; otherwise, if $t > t^*$, no reinsurance will be bought.

260 We notice that it is never optimal to wait for buying reinsurance. That is, it is
261 convenient either to immediately sign the contract, or not to subscribe at all.

262 In particular, at the starting time $t = 0$, given an initial wealth $R_0 > 0$, we have
263 these cases:

264 1. *if $K \geq K^*$, then $\alpha^* = (T, 1)$, that is no reinsurance is purchased;*

265 2. if $K < K^*$, then $\alpha^* = (0, \{\frac{q}{\eta\sigma_0^2}e^{-R(T-s)}\}_{s \in [0, T]})$, that is the optimal choice for the
 266 insurer consists in stipulating the contract at the initial time, selecting the optimal
 267 retention level (as in the pure reinsurance problem).

268 By the expression (6.1) we can show that K^* is increasing with respect to η and σ_0 ,
 269 while it is decreasing with respect to q . More details will be given in the next section by
 270 means of numerical simulations.

271

272 Another relevant result for the reinsurance company is the following.

273 **Proposition 3.** For any fixed cost $K > 0$ there exists $q^* \in (0, +\infty)$ (depending on K) such that

1. if $q > q^*$, then

$$V(t, x) = g(t, x) = \mathbb{E}[e^{-\eta X_T^{t,x}}]$$

274 and $\alpha^* = (T, 1)$, that is no reinsurance is purchased;

2. otherwise

$$V(t, x) = \begin{cases} \bar{V}(t, x - K) = e^{-\eta(x-K)e^{R(T-t)}} e^{\int_t^T \Psi(s) ds} & (t, x) \in [0, t^*] \times \mathbb{R} \\ g(t, x) = \mathbb{E}[e^{-\eta X_T^{t,x}}] = e^{-\eta x e^{R(T-t)}} e^{\int_t^T h(s) ds} & (t, x) \in (t^*, T] \times \mathbb{R}, \end{cases}$$

where $t^* \in (0, T)$ is the unique solution to the equation

$$\eta \left(\frac{q}{R} + K \right) e^{R(T-t)} - \frac{1}{2} \frac{q^2}{\sigma_0^2} (T-t) - \frac{1}{4R} \eta^2 \sigma_0^2 (e^{2R(T-t)} - 1) - \frac{\eta q}{R} = 0,$$

275 and $\alpha^* = (\tau_{t,x}^*, \{\frac{q}{\eta\sigma_0^2}e^{-R(T-s)}\}_{s \in [\tau_{t,x}^*, T]})$, with $\tau_{t,x}^*$ is given in (5.14).

Proof. Following Theorem 4 and its proof, we can write the condition $H(0) \leq 0$ as

$$\frac{T}{2\sigma_0^2} q^2 + (1 - e^{RT}) \frac{\eta}{R} q + \frac{\eta^2 \sigma_0^2}{4R} (e^{2RT} - 1) - \eta K e^{RT} \geq 0.$$

To simplify our computations, let us consider this inequality for any $q \in \mathbb{R}$. The discriminant Δ must be positive, otherwise the existence of $K^* > 0$ in Theorem 4 is not guaranteed anymore. The solutions of the associated equations are

$$q_{1,2} = \frac{\eta\sigma_0^2}{T} \left(\frac{e^{RT} - 1}{R} \pm \sqrt{\Delta} \right), \quad q_1 < q_2.$$

Since

$$q_2 > \frac{\eta\sigma_0^2(e^{RT} - 1)}{RT} > \eta\sigma_0^2,$$

276 only q_1 is relevant because of the condition (3.3). That $q_1 \in (0, +\infty)$ is a consequence of
 277 the existence of $K^* > 0$ in Theorem 4. If q_1 was not positive, then $H(0) > 0$ for any value
 278 of $q > 0$ and this would contradict Theorem 4. Setting $q^* = q_1$ concludes the proof. \square

279 The last result is interesting for the reinsurer. In Section 3 we have already stated
 280 that the condition $q < \eta\sigma_0^2$ (see equation (3.3)) is required in order that the reinsurance
 281 agreement is desirable. In presence of a fixed initial cost, now we know that there exists
 282 a threshold q^* , which is smaller than $\eta\sigma_0^2$, such that the insurer will never subscribe the
 283 contract if $q > q^*$.

284 **Remark 4.** Recalling that $q = \theta\lambda\mu$ (see Section 2, we can give a deeper interpretation of the
 285 previous result. Indeed, we have proven the existence of a maximum safety loading $\theta^* > 0$, which
 286 cannot be exceeded by the reinsurer, otherwise the reinsurance contract will not be subscribed.

287 7. Numerical simulations

288 In this section we use some numerical simulations in order to further investigate
 289 the results obtained in Section 6. Unless otherwise specified, all the simulations are
 290 performed according to the parameters of Table 1 below.

| Parameter | Value |
|------------|-------|
| T | 10 |
| η | 0.5 |
| σ_0 | 0.5 |
| q | 0.1 |
| R | 0.05 |

Table 1: Model parameters.

291 We have previously illustrated how the threshold K^* in equation (6.1) is relevant
 292 for the insurer as well as for the reinsurer. Indeed, K^* turns out to be the maximum
 293 subscription cost that the insurer is willing to pay. The next pictures show how this
 294 threshold is influenced by the model parameters. As expected, if the reinsurer increases
 295 her net profit q , then the fixed cost should decrease, see Figure 1. In practice, recalling
 296 that $q = \theta\lambda\mu$, if the reinsurer increases her safety loading θ , the subscription cost should
 297 be selected from a smaller range $(0, K^*)$. Otherwise, no reinsurance contract will be
 298 stipulated. Let us notice that, according to equation (3.4), any increase of θ implies a
 299 larger retention level as well.

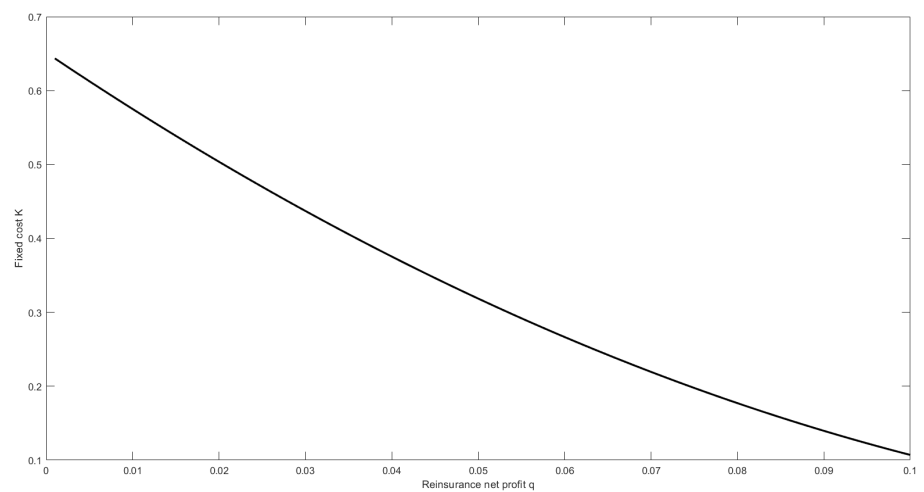


Figure 1. The effect of the reinsurer's net profit q on K^* .

300 As illustrated in Figure 2, when the insurer is more risk averse, she is willing to
 301 pay a higher fixed cost. This result reinforces the practical implications of equation (3.4),
 302 which implies that the more risk averse is the insurer, the larger protection she will buy.

303 Figure 3 shows the effect of the potential losses. When they increase, that is σ_0 is
 304 high, then the insurer is going to pay high fixed cost in order to obtain protection.

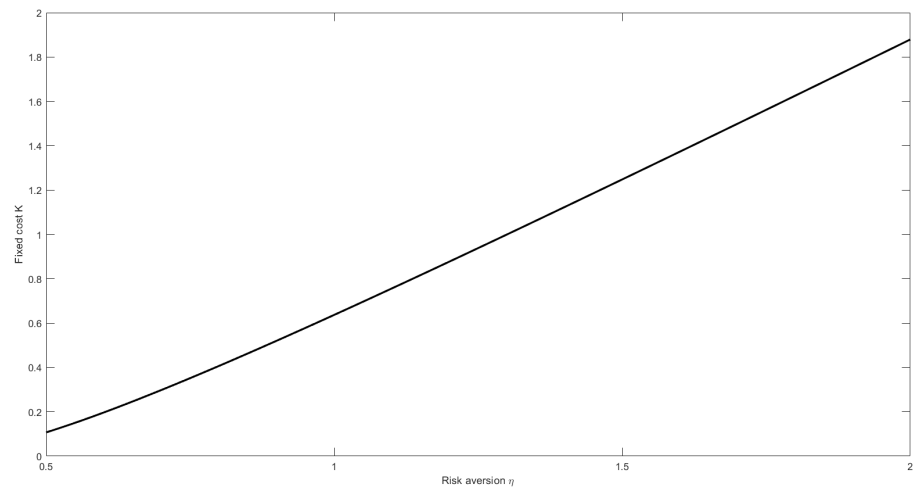


Figure 2. The effect of the risk-aversion parameter η on K^* .

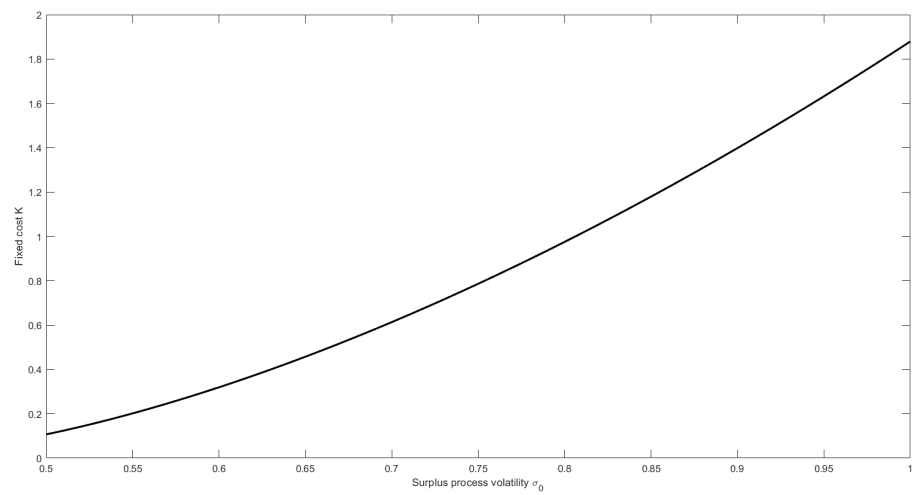


Figure 3. The effect of the volatility parameter σ on K^* .

305 Finally, we can see from Figure 4 that the larger the insurer's time horizon is, the
306 higher the fixed cost will be.

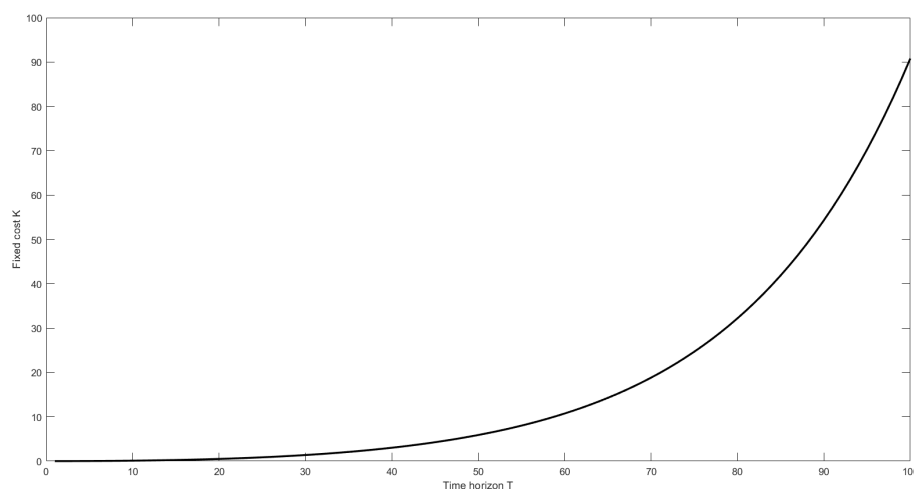


Figure 4. The effect of the time horizon T on K^* .

307 8. Conclusions

308 We have investigated the optimal reinsurance problem under the assumption that
 309 a transaction cost is paid when the agreement is signed. The insurer has to choose
 310 the optimal starting time of the reinsurance contract, as well as the optimal retention
 311 level to be applied from that moment on. We have solved the resulting mixed optimal
 312 control/optimal stopping time problem using a two-steps procedure. We have found
 313 out that the optimal strategy is deterministic (see Theorem 4). Moreover, we have proven
 314 the existence of a maximum fixed cost K^* (see equation (6.1)) that the insurer is willing
 315 to pay. That is, whenever a fixed cost $K > K^*$ is chosen by the reinsurer, the insurer
 316 will retain all her losses. In the last section we have further analyzed how the model
 317 parameters affect that maximum subscription cost.

318 Some future researches could be focused on the study of the optimal reinsurance problem
 319 with fixed cost under either different optimization criteria, or different types of contract.

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