



1 **Analysis of the relationship between yield in cereals and**
2 **remotely sensed fAPAR in the framework of monitoring**
3 **drought impacts in Europe**

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8 **Abstract.** This study analyses the relationship between satellite-measured fAPAR (Fraction
9 of Absorbed Photosynthetically Active Radiation), which are continuously monitored by the
10 European Drought Observatory (EDO) of the EU's Copernicus Emergency Management
11 Service, and crop yield data for cereals, which are collected by Eurostat. Different features of
12 the relationship between annual yield and multiple time series of fAPAR, collected during
13 different periods of the year, were investigated. Two key outcomes of the analysis were the
14 identification of the period from March to October as that when the highest positive
15 correlation between fAPAR and yield is detected in Europe on average, and February to May
16 as the period when most of the negative correlation are observed. While both periods align
17 well with the commonly assumed dynamic of the growing season, spatial differences are also
18 observed across Europe. On the one hand, the Mediterranean regions report the highest
19 correlation values ($r > 0.8$) and the longest continuous periods with positive statistically
20 significant results (up to 7 months), covering most of the growing season. On the other hand,
21 the central European region is characterized by the most limited positive correlation values,
22 with only 2 months or less showing statistically significant results. While marked differences
23 on the overall capability to capture the full dynamic of yield are observed across Europe,
24 fAPAR anomalies seem capable to distinguish between drought and no/drought years in most
25 of the cases if negative yield anomalies are used as a proxy variable for drought impacts.

26

27 **Keywords:** vegetation indices, agricultural drought, EDO, CEMS.

28



29 **1. Introduction**

30 Drought is a multifaceted phenomenon threatening societies, economies and ecosystems in a
31 complex web of cascading effects (UNDRR, 2021). Amongst the major sectors that are
32 impacted by drought, agriculture is still recognized as the most sensitive one (FAO, 2015;
33 FAO et al., 2018), as reflected by the large share of reported impacts for agriculture over the
34 majority of countries and drought events in Europe (Stahl et al., 2016).

35 Most drought monitoring systems recognize the prominent role of agricultural drought,
36 by refining indicators of meteorological drought in order to better account for impacts on
37 vegetation growth (e.g. the Standardized Precipitation-Evapotranspiration Index – or SPEI;
38 Vicente-Serrano et al., 2012), and/or by directly incorporating drought indicators that are
39 based on remotely sensed vegetation indices (WMO and GWP, 2016). In particular, negative
40 deviations from climatological values of satellite measurements of vegetation “greenness” –
41 for example, the standardized anomalies of the fraction of Absorbed Photosynthetically
42 Active Radiation (fAPAR) that are provided by the European and Global Drought
43 Observatories (EDO and GDO, <https://edo.jrc.ec.europa.eu>) – are often adopted as a proxy
44 variable for the adverse effects of drought on vegetation.

45 While such approaches are logically based on the connection between reduced
46 vegetation greenness and diminished plant productivity, it is also well known that droughts
47 occurring during different phenological stages may have different impacts on yield and
48 production (i.e. Barros et al., 2021; Ceglar et al., 2020; Chaves et al., 2002; Demirevska et al.,
49 2009; Stallmann et al., 2020; Zampieri et al., 2017). Consequently, greenness anomalies can
50 be interpreted differently, depending on the development stages of the vegetative cycle when
51 they manifest. Some studies have tried to account for this concept by limiting the analysis to
52 the growing period and excluding data for the plant dormancy phase (e.g. Rojas et al., 2011),
53 by deriving key variation metrics (i.e. amplitude, integral, maximum) from the full growing
54 season (e.g. Kang et al., 2018), or by focusing only on key periods (i.e. a specific month) that
55 have been shown to correlate well with deviations in annual yield for a given study area
56 (Bachmair et al., 2018).

57 Within the framework of the near real-time monitoring of drought events, the task of
58 evaluating and quantifying the actual relevance of an observed anomaly in vegetation
59 greenness is complicated by the need to update continuously the status based on newly
60 acquired data, without the benefit of the full picture of the complete vegetation cycle. This
61 limits the possibility to implement some of the above mentioned approaches as part of



62 operational drought monitoring systems, other than the simple masking of data acquired
63 outside of a pre-defined growing season. An example of an early warning system that
64 accounts for the timing of the observed anomalies is the Anomaly Hot Spots of Agricultural
65 Production (ASAP) decision support system (Rembold et al., 2019), where the seasonal
66 progression (expansion, maturity, senescence) is explicitly considered in determining the
67 warning level.

68 As part of the shift in the drought risk management paradigm from a reactive to a
69 proactive approach, the move from simple hazard indicators to quantitative assessments of
70 risk and impacts is likely to be further integrated within modern early warning systems
71 (UNDRR, 2021). In this regard, independent estimates of actual drought impacts, such as
72 records of yield deviations for different crop types, constitute a valuable reference to assess
73 how the relationship between anomalies in vegetation indices and drought conditions varies in
74 space and throughout the year. This also enables the evaluation of the efficiency of remotely
75 sensed indicators as a proxy for the effect of drought on vegetated land, and the refinement of
76 their use as stress-forcing data for agro-economic models, for the assessment of losses in
77 agriculture due to droughts (García-León et al., 2021).

78 In this context, the primary goal of this study is to analyse to what extent the year-by-
79 year dynamics of yield in Europe can be explained by a regularly updated operational drought
80 indicator, in particular by the fAPAR anomalies produced by EDO. Yield data for cereals,
81 recorded by Eurostat, are used as a general proxy for the impact of droughts on agriculture.
82 The spatio-temporal variations in the relationship between dekadal (i.e. 10-day) fAPAR
83 anomalies and yearly yield deviations can help in identifying the periods of the year when
84 these two quantities are most closely related in the different parts of Europe, thus providing a
85 quantitative basis for improving the assessment of drought impacts in agriculture, with
86 potential benefits both for drought monitoring systems and for agro-economic models in
87 Europe.

88 **2. Material and Methods**

89 **2.1 Eurostat yield dataset**

90 Eurostat, the European Statistical Office, publishes regular reports of statistics on annual
91 crops, including data on production, cultivated area and yield for different crop types, at both
92 national and sub-national aggregation levels (Eurostat, 2020), with the aim of providing a
93 harmonized database of data collected by EU Member States and neighbouring countries.



94 For the purposes of this study, annual data on yield for cereals have been retrieved
95 between 2001 (first full year with available fAPAR data) and 2018 (last available year in the
96 Eurostat database at the time of this study), mostly at the spatial scale of Eurostat’s so-called
97 “NUTS 2” regions (hereafter referred to simply as regions). Only in the case of Germany and
98 the UK data at the NUTS 1 level were used, in order to maximize both data coverage and
99 consistency in region size with the rest of the domain.

100 Because the yield data are to be used for computation of deviations from the long-term
101 average, temporal consistency in the data records is essential. For this reason, records that are
102 flagged by Eurostat as estimated, provisional, unreliable or with a definition that differs due to
103 missing components, were excluded from the analysis.

104 Systematic changes in the annual yield time series were removed by applying a
105 Savitzky–Golay filter to account for advancement in technology and crop management
106 (Tadesse et al., 2015), before standardized anomalies were computed only for those regions
107 with more than 9 years of data (i.e. half of the analyzed period). In this way, 240 regions with
108 valid time series were obtained.

109 **2.2 MODIS fAPAR dataset**

110 The fraction of Absorbed Photosynthetically Active Radiation (fAPAR) is one of the 50
111 Essential Climate Variables recognized by the UN Global Climate Observing System
112 (GCOS), mainly thanks to its direct relationship with primary production
113 (<https://gcos.wmo.int/en/essential-climate-variables/fapar>).

114 fAPAR, and in particular its deviations from historical climatology, constitutes the ideal
115 proxy variable for the effects of drought on vegetated lands (Rossi et al., 2008). In this
116 context, remote sensing images collected by the MODIS (MODerate resolution Imaging
117 Spectroradiometer) sensor represent a unique data source for drought studies, due to the
118 unprecedented longevity of the Terra satellite.

119 In this study, the standard MODIS Terra LAI/fAPAR product (i.e. MOD15A2H,
120 Collection 6) is used (Myneni, 2015), in which global fAPAR maps are derived from the
121 atmospherically corrected Bidirectional Reflectance Distribution Function (BRDF) recorded
122 by MODIS in 7 spectral bands, by solving the three-dimensional radiation transfer process
123 through a look-up-table approach (Knyazikhin et al., 1998; Wang et al., 2001).

124 The standard MODIS product is distributed as 8-day composites (using a maximum
125 composite method) at a spatial resolution of 500-m in $1,200 \times 1,200$ km tiles on a sinusoidal



126 grid. Data include a quality assessment (QA) layer that allows to detect where the simplified
127 back-up algorithm has been used.

128 Datasets of both fAPAR and Δ fAPAR anomalies based on MOD15A2H raw data are
129 regularly produced as part of the European and Global Drought Observatories (EDO and
130 GDO, <https://edo.jrc.ec.europa.eu>) of the EU's Copernicus Emergency Management Service.
131 The operational fAPAR dataset is obtained after a set of pre-processing procedures, including:
132 1) screening of the low-quality data based on the QA flag; 2) spatial aggregation of the data
133 (simple average) at 1-km resolution and re-projection onto a lat/lon regular grid at 0.01°
134 resolution with nearest neighbour resampling; 3) temporal aggregation at dekadal scale (three
135 maps per month: days 1–10, 11–20 and 21–end-of-month) by means of a weighted average of
136 the two closest 8-day images (weight proportional to the overlapping with the dekadal
137 period); and 4) exponential temporal smoothing of the dekadal data (with smoothing
138 parameter equal to 0.5; Brown and Meyer, 1961).

139 Here, the fAPAR anomalies were computed as standardized deviations from the
140 reference period (2001-2018), only if at least 6 years of data were available and only where
141 the long term standard deviation was greater than 0.01 (to exclude areas of low variability,
142 such as deserts or highly stable densely vegetated areas). The reference period of 2001-2018
143 is consistent with the one used for the yield anomalies.

144 **2.3 Analysis strategy**

145 In this study, the analysis of the relationship between the dekad time series of fAPAR
146 anomalies and yearly crop yield is based primarily on computation of the Spearman
147 correlation coefficient (r). In order to carry out the analysis, the two main discrepancies
148 between the two datasets, namely regarding the spatial units (i.e. regions versus cells) and
149 temporal frequency (year versus dekad), must first be considered.

150 Given the focus of the study, the only fAPAR conditions that are relevant are the ones
151 observed over arable land. Therefore, the fAPAR anomaly data were first upscaled to NUTS 2
152 regions as a weighted average of all the 0.01° resolution fAPAR anomaly values within a
153 region, with a weighting factor based on the fraction of each grid-cell classified as arable land
154 according to the latest Corine land cover map (CLC2018, [https://land.copernicus.eu/pan-
155 european/corine-land-cover/clc2018](https://land.copernicus.eu/pan-european/corine-land-cover/clc2018)).

156 Regarding temporal frequency, while fAPAR anomaly data are available throughout the
157 year, similar studies (e.g. Rojas et al., 2011) have focused only on data collected during the
158 growing season. A north-to-south gradient have been observed in the start, the end, and the



159 length of the growing season in Europe, with April-September being a common period all
160 over Europe, but with early start in February and late end in November over many areas
161 (Rötzer and Chmielewski, 2001). Estimations of the growing season directly based on
162 remotely sensed vegetation indices have also highlighted a very early start in the
163 Mediterranean, around October/November of the previous year (i.e., Atzberger et al., 2014),
164 likely related to the effect of infesting weeds on the remote sensing signal. Following these
165 considerations, here we analyse an extended period, testing the relationship between the yield
166 of a particular year and the fAPAR anomalies between the first dekad of October of the
167 preceding year and the end of the current year, for a total of 45 dekadal time series.

168 The set of correlation analyses between each of the 45 dekadal time series of fAPAR
169 anomalies and yearly yield data is used to construct a “correlogram”, which relates the dekad
170 with the corresponding r value (see example in Fig. 1 for the Tuscany region in Italy).
171 Different analyses can be performed on the correlograms, depending on the critical values that
172 are extracted from these plots and on the goal of the analysis. Here, we faced the problem in
173 two different ways: a) detecting periods of similar behaviour and accuracy but variable length;
174 and b) detecting periods of similar length but variable accuracy and behaviour.

175 For these two analyses, we distinguished between two different behaviours in the
176 fAPAR-yield relationship, a direct relationship (i.e. negative anomalies in fAPAR correspond
177 to negative anomalies in yield) and an inverse relationship. The latter may occur when a
178 strong vegetative growth is observed early in the season during drought years, especially in
179 energy-limited conditions (van Hateren et al., 2021). We also distinguished between two
180 levels of accuracy, statistically significant correlations ($p < 0.05$, either positive or negative)
181 and at least different than zero r values (i.e. $|r| > 0.15$).

182 Starting with a minimum length of 2 dekads, up to 990 periods of various length (L ,
183 from 2 up to 45 dekads) can be analyzed for each region, and for each of these periods four
184 main metrics are computed: 1) F_{p+} , the fraction of r values in the period that are positive and
185 statistically significant (i.e. $r > 0$ and $p < 0.05$); 2) F_{p-} , the fraction of r values in the period that
186 are negative and statistically significant (i.e. $r < 0$ and $p < 0.05$); 3) F_+ , the fraction of r values
187 in the period that are at least positive (i.e. $r > 0.15$), and 4) F_- , the fraction of r values in the
188 period that are at least negative (i.e. $r < -0.15$). We can then focus on the longest periods that
189 have homogeneous behaviour and accuracy for a given region, namely homogeneous periods
190 from hereafter (i.e. a period with $F_{p+} = 1$ represents a continuous streak of dekads with all
191 positive and statistical significant r values). In the example reported in Figure 1, the dekads
192 between 23 and 36 (light grey area) are clearly part of the longest period with all positive and



193 statistically significant r values, $L = 14$, while the dark grey area demarks the longest period
194 with $F_{-1} = 1$ ($L = 6$).

195 A further set of analyses is focused on a fixed time window, or a limited range of
196 lengths, which boundary values can be derived from the previous tests. Within these limits, an
197 optimal positive (negative) period for each region can be defined as the period with the
198 maximum (minimum) average r value. Differently from the first group of analyses, these
199 optimal periods have varying F_{p+} and F_{+} (F_{p+} and F_{+}) values that can be used to quantify the
200 robustness of the relationship between fAPAR and yield.

201 Finally, while the analyses based on correlation give an insight on the relationship
202 between fAPAR and yield over the full spectrum of variability, a further test focused only on
203 drought conditions is also performed, given that in the context of drought monitoring it would
204 be sufficient to be able to distinguish drought-affected years from the rest. Here, the total
205 number of cells for each region with fAPAR anomalies < -1 is computed during the drought
206 years (yield anomalies < -1), and it is compared with the same during no-drought years (yield
207 anomalies ≥ -1). The assumption here is that the ratio of these two quantities should be greater
208 than one in the case of a direct relationship.

209 **3. Results and Discussion**

210 **3.1 Yield anomalies as a proxy variable for drought impacts**

211 Since standardized yield anomalies for cereals, as described in Section 2.1, are here used as a
212 proxy variable for the effects of droughts on vegetated land, a preliminary analysis was
213 performed to evaluate the reliability of this assumption for the derived dataset.

214 Figure 2 depicts the temporal evolution of yearly yield deviations, highlighting some
215 clear spatial patterns of significantly negative anomalies (i.e. yield anomaly < -1). Following a
216 review of the scientific literature for past drought events, it is possible to associate a
217 documented main drought event to each of these large clusters, as summarized in Table 1.
218 Seven main droughts are reported, ranging from the well-known drought in central Europe of
219 2003 (Rebetz et al., 2006) to the central-north European drought of 2018 (Buras et al., 2020;
220 Toreti et al., 2019).

221 The existence of a cause-effect relationship between the largest spatial patterns
222 observed in negative yield anomalies and the listed major drought events is further supported
223 by the study of Spinoni et al. (2015), which categorized the listed events (except the last two,
224 which occurred after that study) as being among the most severe in Europe according to
225 meteorological drought indices.



226 For each of the drought events listed in Table 1, specific independent scientific
227 references are also provided, which include details on the evolution of the meteorological
228 conditions, and the potential impacts on agriculture. Overall, analyses of these data tend to
229 support the underlying assumption that yield anomalies can be used as a reliable independent
230 estimate of drought impacts on vegetation, in conformance with the conclusions of other
231 studies at regional level in Europe (Bachmair et al., 2018; Potopová et al., 2015), or for other
232 parts of the world (e.g. Yang et al., 2020).

233 3.2 Detection of the homogeneous periods in the fAPAR-yield relationship

234 While many studies focused on the local maximum r value to detect when and where
235 fAPAR and annual yield anomalies best correlate, isolated peak values may alter the
236 perception of the robustness of fAPAR as a proxy variable of yield. In the context of an
237 operational drought monitoring system, where continuous estimates should be provided rather
238 than “one shot” predictions, information on longer homogeneous time periods are more
239 valuable.

240 Focusing first on the positive r values, we analysed the periods with only statistically
241 significant values ($F_{p+} = 1$), or only at least positive values ($F_+ = 1$). The maps in Fig. 3
242 reports the local maximum lengths corresponding to these two quantities, namely positive
243 homogeneous periods. Both of these maps show generally longer homogeneous periods in
244 southern Europe, with the largest values observed for some Mediterranean regions (e.g. most
245 of Spain, Cyprus, Sicily, Apulia and the Aegean/Mediterranean Turkey), and the smallest
246 values (or no homogeneous period at all) mostly located in Central Europe (i.e. Germany,
247 Poland and north-eastern France). On average, the maximum length of the periods with F_{p+}
248 $= 1$ is limited in most of the cases (5.5 ± 4.3 dek, almost 2 months), whereas the values more
249 than double in the case of $F_+ = 1$ (13.0 ± 8.3 dek, more than 4 months).

250 Generally, almost all the maximum r values in the correlograms are obtained in the
251 dekads between mid-February and mid-September, which is expected since this period aligns
252 well with what is commonly considered the growing season in Europe (Atzberger et al., 2014;
253 Rötzer and Chmielewski, 2001). Nonetheless, a large variability in the length of both positive
254 homogeneous periods is observed, with southern and central Europe confirmed to be not only
255 the areas with highest and lowest r values, respectively, but also the areas with the longest (i.e.
256 4-7 months) and shortest (up to 2 months) periods with consecutive statistically significant
257 positive correlations.



258 In order to evaluate synthetically the temporal location of these homogeneous periods,
259 we analyzed which dekads each of them covers, and computed for every dekad the fraction of
260 NUTS 2 regions (out of 240) that includes that particular dekad in the homogeneous period
261 (Fig. 4). For example, dekad 27 (i.e. the first dekad of July starting from the beginning of
262 October of the previous year) is part of the maximum homogeneous period in about 20% and
263 50% of the regions, for F_{p+} and F_+ , respectively. It is worth noting that about 21% of the
264 NUTS 2 regions do not have a period (minimum 2 consecutive dekads) with $F_{p+} = 1$.

265 It is possible to observed two “flexing points” in each of the two time series in Fig. 4:
266 both around the values 0.1 for F_{p+} and around 0.2 for F_+ . Starting from these values, we can
267 detect two optimal homogeneous periods: from end-of-April to mid-October (6 months) for
268 F_{p+} , and from March to early-November (8 months) for F_+ .

269 Moving to the negative correlation values, two maps analogous to the ones in Fig. 3 are
270 reported in Fig. 5 for F_{p-} (panel a) and F_- (panel b). These two maps show how the longest
271 negative homogeneous periods are in general shorter than the ones for positive correlations,
272 with an average value of 3.0 ± 1.6 dekads for F_{p-} and 7.0 ± 3.9 for F_- . The lack of statistically
273 significant negative r values is especially evident, with almost 50% of the regions having no
274 homogeneous periods with $F_{p-} = 1$. The map for F_- (Fig. 5b) allows for some additional
275 considerations on the spatial distribution, with moderate maximum lengths (around 9 dekads)
276 in most of western and central Europe, and some high values (higher than 15 dekads) in some
277 regions of southern Europe.

278 In terms of temporal distribution, the histograms on Fig. 6 depict the fraction of NUTS
279 2 regions that includes that particular dekad in the negative homogeneous periods. Overall,
280 the fraction values are lower than the ones observed for the positive periods (see Fig. 4), with
281 two distinguishable peak periods in the F_- values, the first in early season (February-May) and
282 the second at the end of the season (October-December).

283 Most of the homogeneous periods early in the season correspond to regions in western
284 and southern Europe, and the late season periods are mostly located in central and northern
285 Europe. In the framework of drought monitoring, the first can be potentially exploited as early
286 warning signals of subsequent reduction in fAPAR due to drought (as seen in the positive
287 homogenous periods), while the second mostly occur very close to the harvesting season.

288 3.3 Performance for a fixed time-window

289 A clear outcome of the previous analyses is that the length of the homogeneous periods
290 with negative correlations is limited compared to the positive correlations, and mostly useful



291 for drought monitoring only early in the growing season. Therefore, we focus only on the
292 positive correlation values for the successive analyses. The two lengths (6 and 8 months)
293 derived from the data depicted in Fig. 4 are used as the minimum and maximum boundary
294 values to find the local optimal period for each region (see Section 2.3).

295 The results of this bounded analysis of the local optimal period are shown in Fig. 7,
296 where the starting dekad (d_i , panel a) and ending dekad (d_e , panel b) of the optimal period are
297 depicted for every region. Fig. 7a shows a general pattern of an early start in Central Europe
298 (i.e. February/March), and in few southern regions of the Mediterranean, and a late start (i.e.
299 May/June) in most of southern and western Europe. This late start is of course in line with the
300 previously observed negative correlations in February/May over the same regions.
301 Analogously, Fig. 7b shows that the end of the optimal period occurs mostly around
302 October/November in both southern and western Europe, and August/September in central
303 Europe, with then mostly negative correlations in central and north Europe occurring after this
304 period.

305 Given that these optimal periods have been derived based on the average r values in the
306 6 to 8-month period, both F_{p+} and F_+ can assume any values between 0 and 1. For this reason,
307 we classified each region based on the combined values of these two metrics, as represented
308 by the legend included in Fig. 8. In this map, the green areas show a good capability to
309 reproduce the dynamic of yield deviation for the whole optimal period, with the regions in
310 dark green having the overall best performance (over half of dekads with statistically
311 significant r values and more than 2/3 with at least positive values). Conversely, the red
312 regions show a poor capability of the fAPAR anomalies to capture the yield dynamic, with the
313 dark red regions having less than 1/10 of statistically significant values (i.e. less than a month)
314 and less than 1/3 of positive correlations during the optimal period.

315 Overall, slightly more than half (i.e. 55.8%) of the study regions are classified in one of
316 the green classes, with a predominance of these regions in Mediterranean and south-eastern
317 Europe. The rest of the study area is almost equally split between regions with average
318 performance (yellow class, 23.3%), and poor performance (red classes, 20.9%). Among the
319 red classes, the majority of the regions fall in the category with intermediate F_+ values ($1/3 <$
320 $F_+ < 2/3$) but low statistical significance ($F_{p+} < 1/10$). Most of these regions are located in
321 central Europe, between northern France, the United Kingdom, Germany and Poland.

322 Spain stands out as having particularly robust performances, even among the generally
323 good performing Mediterranean area. While the start and end of the optimal period varies
324 across the area (March to May, and September to November, respectively), the results are



325 consistently in the best class (dark green in Fig. 8). Among the Mediterranean countries, some
326 mixed results can be observed in Italy and Greece.

327 **3.4 Drought vs. no-drought years**

328 The previous analyses show a noticeable difference in the capability of fAPAR
329 anomalies to capture the full range of variability of yield anomalies across Europe, as
330 quantified by the results on the optimal periods summarize in Fig. 8. For the same optimal
331 periods, the number of fAPAR anomalies < -1 where cumulated for drought and no-drought
332 years, separately, and the ratio between these two quantities is depicted in Fig. 9.

333 Overall, values greater than 1 are observed over most of Europe in Fig. 9, suggesting a
334 good capability of fAPAR anomalies to detect the effects of drought on annual yield. While
335 the ratio is only slightly higher than one in some regions where the previous analyses
336 highlight poor performances (i.e. UK and France), effects of the main drought events on yield
337 are still well-captured.

338 Finally, the plot in Fig. 10 show a comparison between the ratio computed on the
339 optimal period (grey area) and the one computed on the full year (all 36 dekads, black area).
340 This plot show an overall increase in the ratio when only the dekads in the optimal period are
341 considered, which translate in a better ability to discriminate between drought and no/drought
342 years compared to simply account for all the anomalies observed across the full year.

343 **4. Discussion**

344 The results reported in the previous section are based on the assumption that anomalies
345 of cereal yields are a good proxy variable for drought impacts, as demonstrated for example
346 by Brás et al. (2021) who quantified an approximately 9% reduction in European cereal yields
347 due to historical droughts (1961-2018), with an increasing intensity in more recent years. The
348 spatial patterns in negative yield anomalies used in this study, and the cross comparison with
349 documented past drought events, confirm this general assumption.

350 The common element for all the performed analyses is the independent use of each
351 dekadal fAPAR time series, since this is how data are commonly used in operational
352 monitoring systems. While different results may be achieved by using metrics based on the
353 full growing season (e.g. Kang et al., 2018), such analyses are not easily transferable to a
354 near-real time monitoring framework. Overall, the correlation coefficients computed using
355 fAPAR collected during multiple dekads suggests a predominance of positive values over all
356 regions. This is in line with the expected direct relationship between fAPAR and yield during
357 the core growing season, as well as with most of the past studies which focused primarily on



358 the positive correlation. Indeed, most of the maximum values of correlation seems to be
359 located within the conventional growing season, and the south-north gradient observed in
360 both of the positive homogeneous period maps (Fig. 3) is in broad agreement with the
361 expected increasing gradient in growing season length observed over Europe (Rötzer and
362 Chmielewski, 2001). However, while studies on satellite-derived phenology have detected
363 growing season lengths ranging from 5 to 9 months (Rötzer and Chmielewski, 2001), the
364 average length of the periods with positive and statistically significant correlations seems to
365 be shorter.

366 Consistently high positive correlation values are obtained over most of Spain, in line
367 with a recent study over the region (García-León et al., 2019), which reported good
368 performances of the satellite-based Vegetation Condition Index (VCI) for different type of
369 cereals, especially for winter wheat and barley. Over central Italy, Todisco et al. (2008)
370 observed good correlation between yield in sunflower and sorghum with common drought
371 indices (Standardized Precipitation Index, SPI, and Soil Moisture Severity Index), with a
372 maximum correlation around weeks 27-29 of the growing season (i.e. July) and statistically
373 significant values for periods ranging from 2 to 4 months. Similar timing, but with a slightly
374 shorter optimal length, has been observed in our analysis for the same area.

375 For Germany, Bachmair et al. (2018) found significant correlation values between VCI
376 and Vegetation Health Index (VHI) anomalies in the month of August, and yield deviations
377 for maize, that are comparable with the maximum values observed for western Germany in
378 our study. A mix of high correlation and missing data is reported in that study for eastern
379 Germany, where our results are statistically significant only for a very limited period. These
380 differences may be explained by the focus on specific crop types, as the same authors also
381 highlight how their results varied for the different crops.

382 Similar to our results, Labudová et al. (2017) found significant correlation with SPI and
383 Standardized Precipitation Evapotranspiration Index (SPEI) in the Danubian lowlands only
384 for summer months, or for a very limited time (i.e. June) in the Eastern Slovak lowlands. For
385 these regions, the values of the maximum homogeneous period with $F_{p+}=1$ ranged between 3
386 and 9 dekads as shown in Fig. 3.

387 The presence of limited periods with consecutive negative correlations early in the
388 growing season may be related to the lagged response of vegetation to water deficits (Crow et
389 al., 2012), as well as to the limited immediate effect of water deficit during energy-limited
390 periods (Zscheischler et al., 2015). This inverse relationship is currently under-explored in
391 drought monitoring systems, which mostly focus on the direct relationship, and it may have



392 an interesting role as an early warning tool under specific conditions. However, the results
393 obtained in this study suggest a limited temporal extension and statistical robustness of the
394 periods with inverse relationships, which usually are followed by much longer and robust
395 periods of direct relationship.

396 The late start of the optimal period in many regions of the Mediterranean and western
397 Europe, compared to the rest of the domain, is associate to the presence of these periods of
398 inverse relationship early in the growing season. Given the particular climate of the
399 Mediterranean region, and the key role of the dry and hot summer months in propagating the
400 droughts in the area, a lagged response in vegetation is expected. In contrast, Central Europe
401 is characterized by an early start of the optimal period (March to August) that seems to
402 precede the expected growing season (June to October). For central Europe, Potopová et al.
403 (2015) found high yield-drought correlation for cereals (better than other crops) over Czech
404 Republic between April-June, a result in line with our findings. The late start (April/May) in
405 the northern regions of Scandinavia compared to central Europe, is mostly explained by the
406 lack of reliable fAPAR data earlier in the year, due to low sun angles.

407 Focusing on the optimal period, mixed performances are obtained in Italy, with low
408 agreement particularly in Sardinia and regions along the Apennine mountains. Although
409 García-León et al. (2021) found a positive relationship between annual-cumulated fAPAR
410 anomalies and yield for most main crop types, the aggregation of the results at national scale
411 does not allow the detection of differences among regions. Given the complex morphology of
412 those regions, potential unreliability in the fAPAR estimates may be a possible cause for the
413 poor performances. Complex morphology can also be the reason for poor results over few
414 other Mediterranean areas, such as Greece.

415 One possible contributing factor underlying the spatial differences in the retrieved
416 optimal periods, is the effects of different predominant cereal types that are cultivated locally.
417 This is supported by other studies that have demonstrated different responses for different
418 crop types (García-León et al., 2021; Labudová et al., 2017). While applying the analysis to
419 different plant types may be useful to understand better the relationship between drought
420 conditions and yield for each specific crop, the results of this study for all cereals provide
421 valuable experimental information that can be more easily ingested into an operational
422 drought monitoring system, which specializes not only in agricultural drought impacts.

423

424



425 5. Summary and Conclusions

426 In this study, records of annual crop yield data for cereals, which are collected by Eurostat,
427 were used to evaluate the capability of satellite-derived fAPAR time series data to capture the
428 effect of drought events on crop production for different periods of the year and growth stages
429 of vegetation.

430 Overall, the analysis of the correlograms computed by plotting anomalies of dekadal
431 fAPAR values against yearly yield deviations, was used for three main purposes:

- 432 ▪ Investigation on continuous streaks of dekads with homogeneous behaviour (direct vs.
433 inverse) and agreement (i.e. statistical significance) but with different temporal length.
- 434 ▪ Investigation of fixed length (6 to 8 months) optimal periods, defined as function of
435 the maximum average r within the given range of lengths.
- 436 ▪ Evaluation of the capability of fAPAR anomalies during the optimal periods to
437 discriminate between drought and no-drought years.

438 The analyses confirm the March to October period as the most relevant to positively
439 correlate anomalies of fAPAR and crop yield, being the period when most of the highest
440 values of correlation are observed, and when most of the continuous periods with statistically
441 significant and positive r values are located. There is a generally good agreement comparing
442 these findings with both the duration and temporal location of the commonly defined growing
443 seasons in Europe. While some periods with consistent negative correlations are also observed
444 between February and May, these are generally limited in length to be considered as primary
445 source of information to reproduce yield dynamics, but they have potential as valuable early
446 warning information.

447 The average growing period in Europe is usually characterized by a marked south-to-
448 north gradient, which is also observed in our analysis of the 6 to 8-month optimal periods
449 based on average r values. Some clear spatial patterns emerge in this analysis, such as the
450 early start in most of central Europe and the southern Mediterranean, and the late start in
451 southern and western Europe. While these spatial patterns do not exactly match commonly
452 observed satellite-derived growing seasons, they provide potentially valuable information that
453 can be incorporated into operational drought monitoring systems.

454 Another key output of the study is the generally good correlation between fAPAR
455 anomalies and crop yield anomalies over most of the Mediterranean regions and across the
456 full range of variability of yield data. Given the well documented high vulnerability of this
457 region to drought and the increasing threat posed by climate change (Cammalleri et al., 2020;



458 Dubrovský et al., 2014), the ability to link satellite-observed fAPAR anomalies with actual
459 impacts in agriculture is a promising new development that merits further exploration.

460 This study also highlighted the overall limited correlation, outside of very short time
461 periods, between fAPAR and yield over most of the NUTS 2 regions in central Europe.
462 Further analyses may be needed to better understand the reason behind this result. In this
463 context, a recent study by Beillouin et al. (2020) has demonstrated how simple climate
464 variables (i.e. high temperature and low precipitation) can explain much of the yield
465 variability in central Europe, in contrast with the situation in southern Europe. It is important
466 to further remark that even over these regions where the overall performance is limited,
467 fAPAR anomalies are still successful in discriminating between drought and no-drought
468 conditions, which is a major requirement in drought monitoring systems.

469

470

471 **Data availability:** The fAPAR dataset used in this study can be retrieved from the JRC Data
472 catalogue (EDO, 2021).

473

474 **Author contribution:** CC designed the experiments with inputs from AT and NMC. CC
475 developed the codes and performed the analyses. CC prepared the manuscript with
476 contributions and revisions from all co-authors.

477

478 **Competing interests:** The authors declare that they have no conflict of interest.



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- 653



654 **Tables**

655

656 **Table 1.** Main European drought events between 2001 and 2018 according to the maps
657 reported in Fig. 2, and corresponding references in the scientific literature.

<i>Year of drought event</i>	<i>Area affected</i>	<i>Reference</i>
2003	Central Europe	Rebetez et al. (2006)
2005	Iberia Peninsula	García-Herrera et al. (2007)
2006	North-Eastern Europe	Valiukas (2015); Somorowska (2016)
2007	Eastern Europe	Bogdan et al. (2008); Sima et al. (2015)
2012	Eastern Europe	Sima et al. (2015)
2017	Southern Europe	García-Herrera et al. (2019)
2018	Central-Northern Europe	Buras et al. (2020); Toreti et al. (2019)

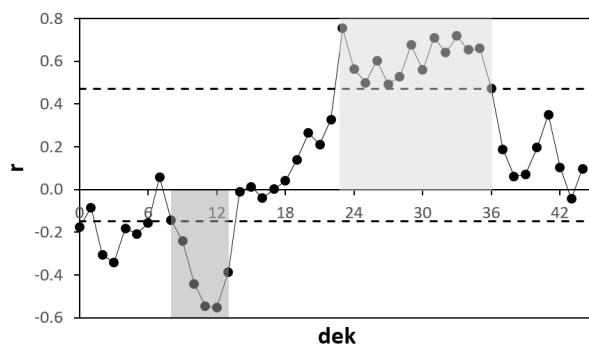
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659



660 **Figures**

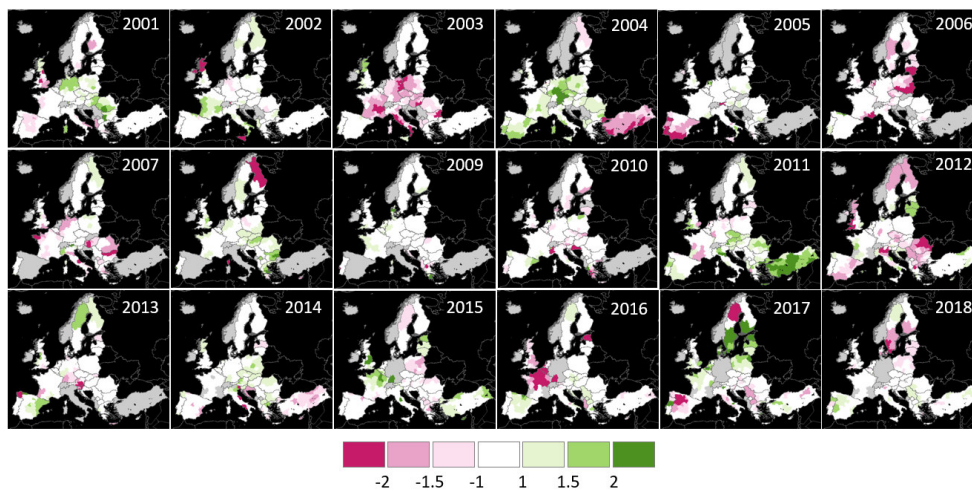
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662

663 **Fig. 1.** Example of correlogram for one NUTS 2 region in Italy (ITI1, Tuscany). Each value
664 represents the Spearman correlation coefficient between the fAPAR anomaly time series of a
665 specific dekad and the yearly yield anomalies. The two horizontal dashed lines represent the
666 positive statistical significant value at $p = 0.05$ and the minimum negative threshold ($r = -$
667 0.15). Dekads are defined starting from the first dekad of October of the previous year (e.g.
668 dek = 23 refers to the last dekad of May of the current year).

669



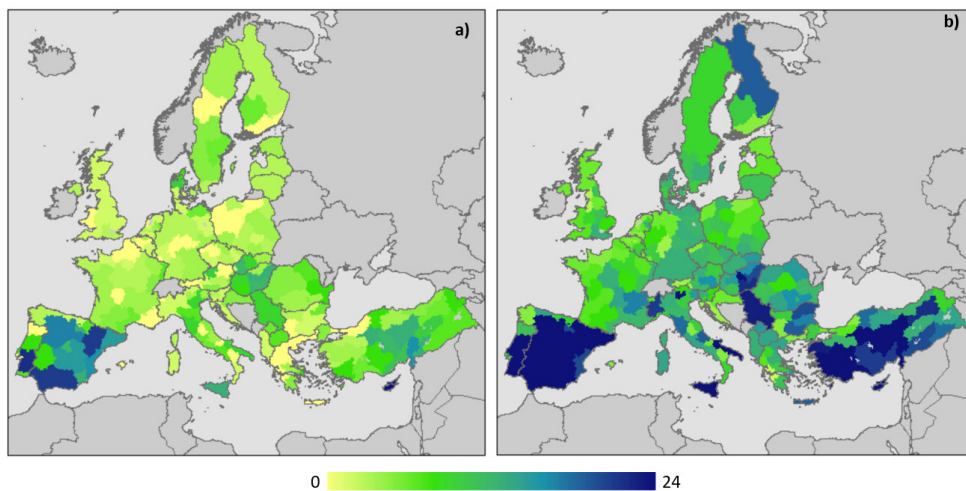
670

671 **Fig. 2.** Spatial distribution of annual standardized yield anomalies for the period 2001-2018.

672 Anomalies are mapped at NUTS 2 level, with the exception of the areas detailed in Section

673 2.1. Data in grey are missing.

674

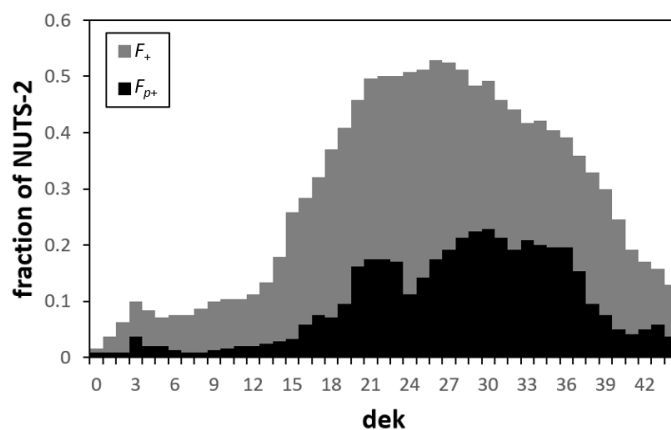


675

676 **Fig. 3.** Spatial distribution of the length (in dekads) of the longest period with $F_{p+} = 1$ (panel

677 a) and $F_+ = 1$ (panel b).

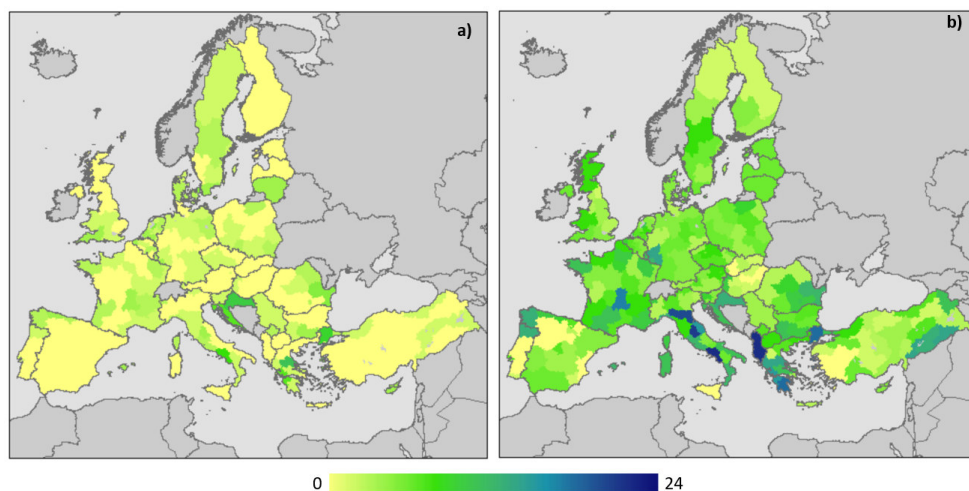
678



679

680 **Fig. 4.** Fraction of NUTS 2 regions for which each dekad is included in the longest
681 homogeneous period with $F_{p+} = 1$ (black) or $F_+ = 1$ (grey).

682

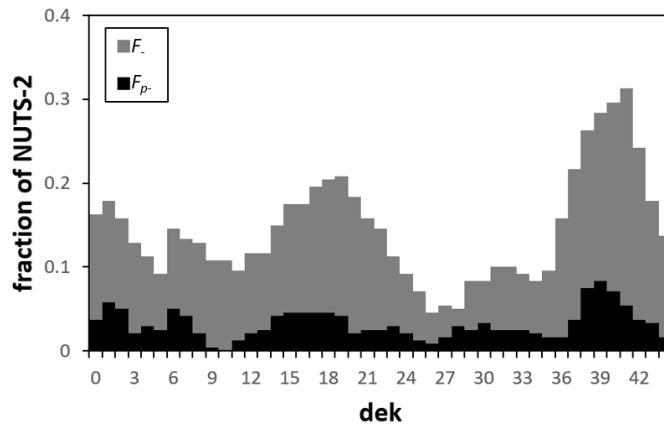


683

684 **Fig. 5.** Spatial distribution of the length (in dekads) of the longest period with $F_p = 1$ (panel

685 a) and $F_p = 1$ (panel b).

686

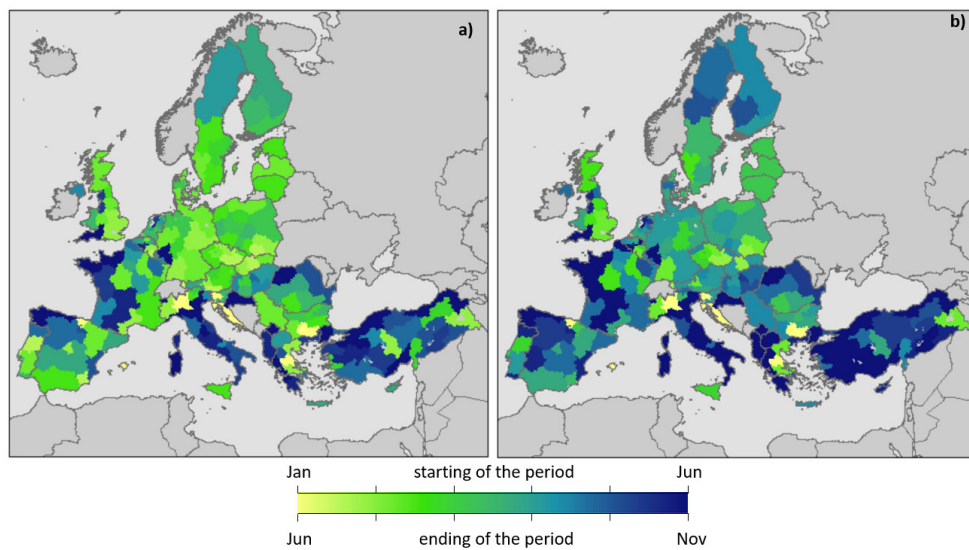


687

688 **Fig. 6.** Fraction of NUTS 2 regions for which each dekad is included in the longest

689 homogeneous period with $F_p = 1$ (black) or $F = 1$ (grey).

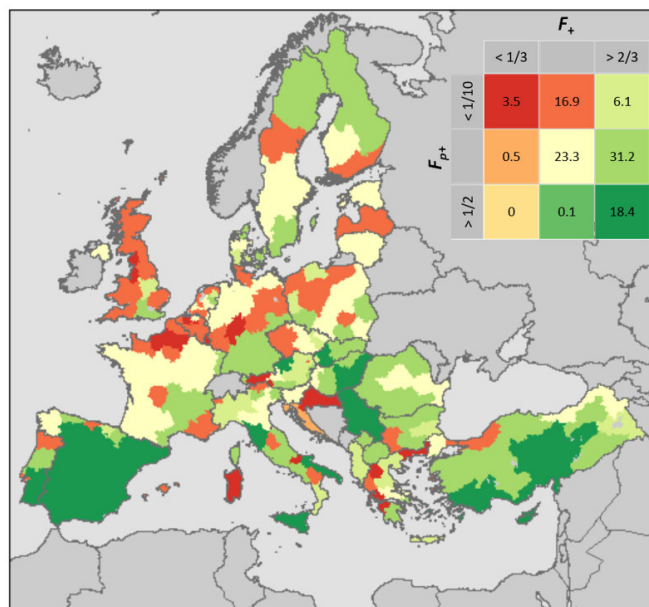
690



691

692 **Fig. 7.** Spatial distribution of (a) the starting dekad, and (b) the ending dekad, of the local
693 optimal period based on the average correlation and bounded by a length from 6 to 8 months.

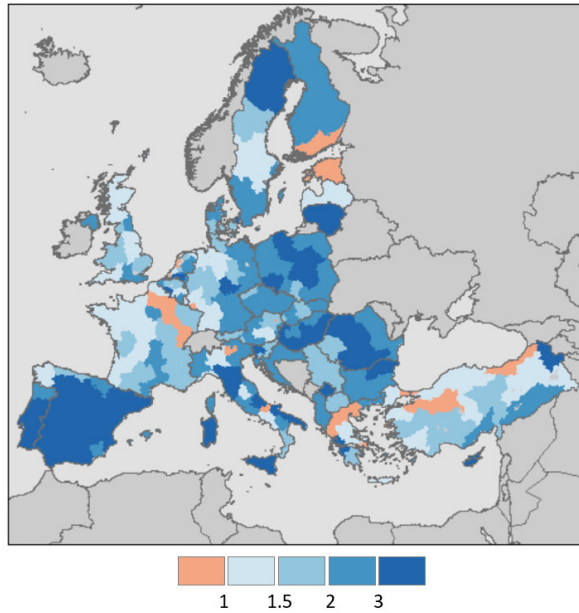
694



695

696 **Fig. 8.** Synthetic representation of the performance of dekadal fAPAR anomalies in
697 reproducing the yearly yield variations during the local optimal period. The inserted legend
698 shows the values of F_{p+} and F_+ for each category, with the numbers inside each square
699 representing the percentage (%) of the total NUTS 2 regions (out of 240) that falls under each
700 category.

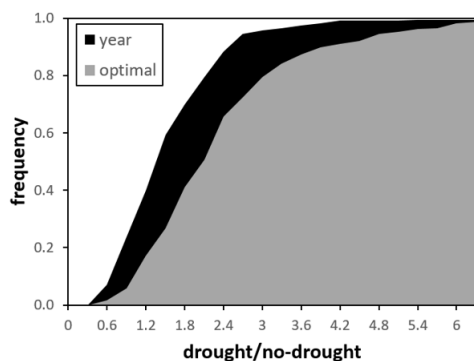
701



702

703 **Fig. 9.** Spatial distribution of the ratio between the number of fAPAR anomalies < -1 in the
704 optimal period (see section 3.3) during drought years (yield anomalies < -1) and no-drought
705 years (yield anomalies ≥ -1).

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708 **Fig. 10.** Cumulated frequency of: i) the ratio between the number of fAPAR anomalies < -1 in
709 the optimal period (see section 3.3) during drought years (yield anomalies < -1) and no-
710 drought years (yield anomalies \geq -1) (optimal, grey area); and ii) the ratio between the number
711 of fAPAR anomalies < -1 in the full year (36 dekads) during drought and no-drought years
712 (year, black area).